Introduction to Probabilistic Reasoning

2. From uncertain events to uncertain numbers

Giulio D'Agostini

Dipartimento di Fisica Università di Roma La Sapienza Preambolo

Ched'è la statistica? (Trilussa)

\Rightarrow preambolo a lezioni al CERN

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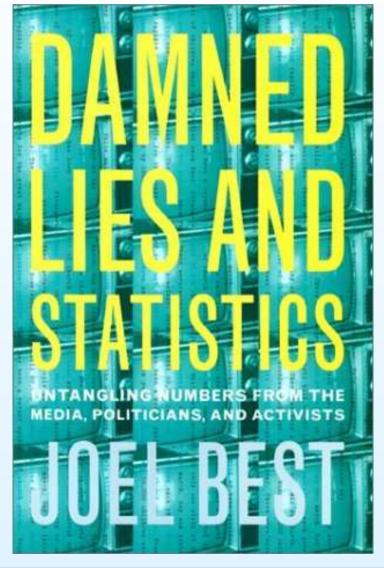
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 - 2. "Tell the truth" \leftrightarrow "to lie"? \Rightarrow Not fair, though

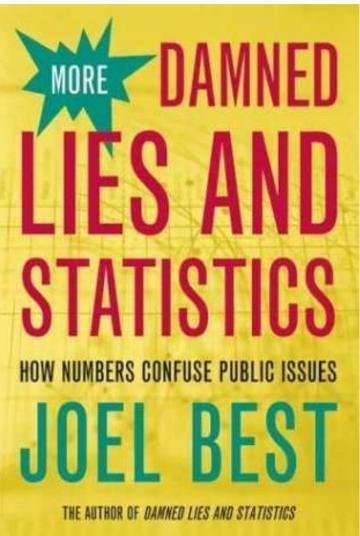
"There are three kinds of lies: lies, damn lies, and statistics" (Benjamin Disraeli/Mark Twain) Damned lies and statistics

Well known subject

Damned lies and statistics

Well known subject, especially in marketing and politics





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Usually several things:

- *descriptive* statistics [e.g. Webster's (Kdict)]
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- Probability theory
- Inference \Rightarrow primary interest to physicists

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Inference: learning about theoretical objects from experimental observations (see later)

Descriptive statistics Little to comment, apart that the process of summarizing 'a State' in a few numbers, in a diagram or in a table causes an enormous loss of detailed information, and this might lead to misunderstandings or even 'lies'.

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- Probability theory Essentially OK, if we only consider the mathematical apparatus.
- Inference Messy:
 - Traditionally, a collection of *ad hoc* prescriptions
 ... accepted more by authority than by full awareness of what they mean
 - ⇒ The physicist is confused[†] between good sense and statistics education

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Inference Do better?

- Much improvement is gained if inference is grounded on probability theory
- Summaries of descriptive statistics can be used in those cases in which statistical sufficiency holds

 (e.g. when we use the sample arithmetic average and standard deviation, instead of the n data points)

Torniamo a noi

Punto della situazione

Outline of first meeting

- Brainstorm on 'standard' teaching of data analysis methods. Problems with confidence intervals and p-values
- Uncertainty, probability, decision.
- Causes \longleftrightarrow Effects

"The essential problem of the experimental method" (Poincaré).

- A toy model and its physics analogy: the six box game "Probability is either referred to real cases or it is nothing" (de Finetti).
- Probabilistic approach, but What is probability?
- Basic rules of probability and Bayes rule.
- Bayesian inference and its graphical representation: ⇒ Bayesian networks
- Let us play for a while with the toy

Summary on probabilistic approach

- Probability means how much we believe something
- Probability depends on available information \rightarrow subjective
- Probability values obey the following basic rules
 - 1. $0 \leq P(A \mid I) \leq 1$
 - 2. $P(\Omega \mid \mathbf{I}) = 1$
 - 3. $P(A \cup B \mid \mathbf{I}) = P(A \mid \mathbf{I}) + P(B \mid \mathbf{I}) \quad [\text{ if } P(A \cap B \mid \mathbf{I}) = \emptyset]$
 - 4. $P(A \cap B \mid \mathbf{I}) = P(A \mid B, \mathbf{I}) \cdot P(B \mid \mathbf{I}) = P(B \mid A, \mathbf{I}) \cdot P(A \mid \mathbf{I})$
- All the rest by logic
- \rightarrow And, please, be coherent!

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 \Rightarrow more comments on $P(E \mid I) \rightarrow$

Three boxes 'paradox'

- 1. The guest and two contestants
- 2. The guest and one contestant

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Nr. 2 \rightarrow Monty Hall problem

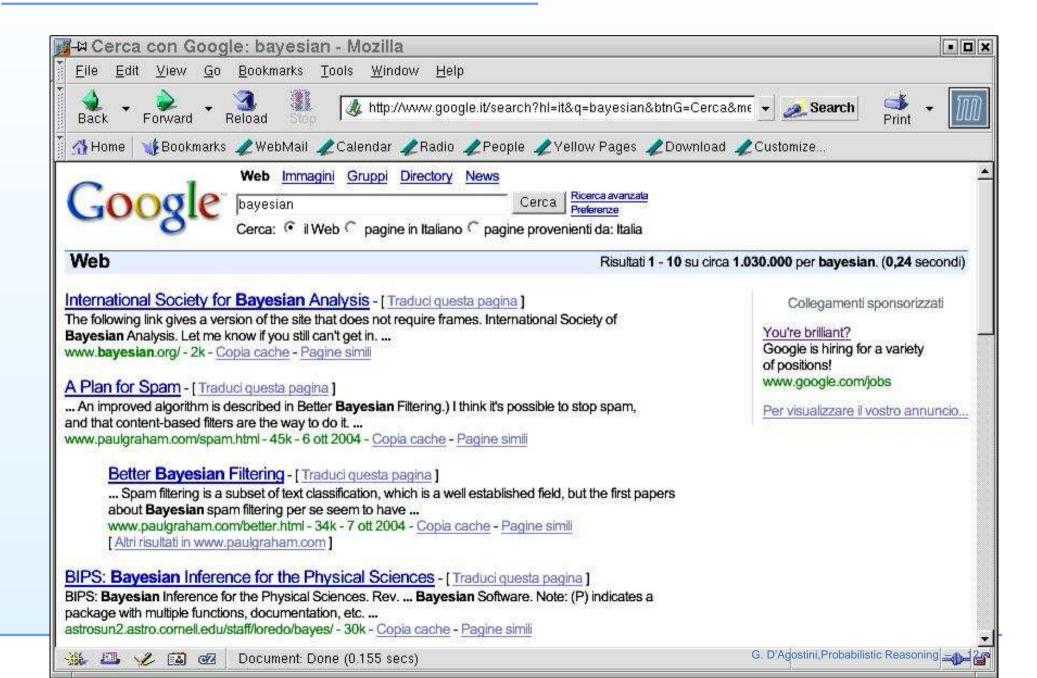
Conclusions on intro to probabilistic reasoning

- Subjective probability recovers intuitive idea of probability.
- Nothing negative in the adjective 'subjective'. Just recognize, honestly, that probability depends on the status of knowledge, different from person to person.
- Most general concept of probability that can be applied to a large variety of cases.
- The adjective Bayesian comes from the intense use of Bayes' theorem to update probability once new data are acquired.
- Subjective probability is foundamental in decision issues, if you want to base decision on the probability of different events, together with the gain of each of them.
- Bayesian networks are powerful conceptuals/mathematical/ software tools to handle complex problems with variables related by probabilistic links.

Are Bayesians 'smart' and 'brilliant'?

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Are Bayesians 'smart' and 'brilliant'?



Further comments on first meeting

The three models example

Choose among H_1 , H_2 and H_3 having observed x = 3:

In case of 'likelihoods' given by pdf's, the same formulae apply: " $P(\text{data} | H_j)$ " \longleftrightarrow " $f(\text{data} | H_j)$ ".

$$\begin{array}{c} 0.5 \\ f(x | H_{1}) \\ 0.4 \\ 0.3 \\ H_{1} \\ 0.2 \\ -2 \\ 2 \\ 4 \\ 6 \\ 8 \\ 10 \end{array}$$

 $BF_{j,k} = \frac{f(x=3 \mid H_j)}{f(x=3 \mid H_k)}$

 $BF_{2,1} = 18$, $BF_{3,1} = 25$ and $BF_{3,2} = 1.4 \rightarrow$ data favor model H_3 (as we can see from figure!), but if we want to state how much we believe to each model we need to 'filter' them with priors.

Assuming the three models initially equally likely, we get final probabilities of 2.3%, 41% and 57% for the three models.

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- But until you don't have an alternative and credible model to explain the data, there is little to say about the "chance that the data come from the model", unless the data are really impossible.
- Why do frequentistic test often work? \rightarrow Think about... (Just by chance – no logical necessity)

Exercises and discussions

- Continue with six box problem [\rightarrow AJP 67 (1999) 1260] \rightarrow Slides
- <u>Home work 1</u>: AIDS problem $\rightarrow P(HIV | Pos)$?

 $P(\mathsf{Pos} | \mathsf{HIV}) = 100\%$ $P(\mathsf{Pos} | \overline{\mathsf{HIV}}) = 0.2\%$ $P(\mathsf{Neg} | \overline{\mathsf{HIV}}) = 99.8\%$

• <u>Home work 2</u>: Particle identification:

A particle detector has a μ identification efficiency of 95%, and a probability of identifying a π as a μ of 2%. If a particle is identified as a μ , then a trigger is fired. Knowing that the particle beam is a mixture of $90\%\pi$ and $10\%\mu$, what is the probability that a trigger is really fired by a μ ? What is the signal-to-noise (S/N) ratio?

What was the mistake of people saying $P(\overline{\text{HIV}} | \text{Pos}) = 0.2$?

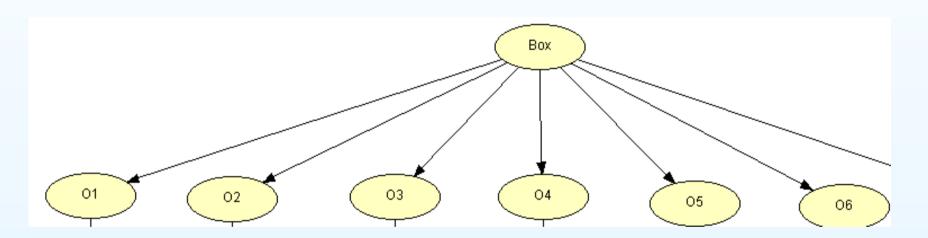
We can easily check that this is due to have set $\frac{P_{\circ}(HIV)}{P_{\circ}(HIV)} = 1$, that, hopefully, does not apply for a randomly selected Italian.

- This is typical in arbitrary inversions, and often also in frequentistic prescriptions that are used by the practitioners to form their confidence on something:
- $\rightarrow\,$ "absence of priors" means in most times uniform priors over the all possible hypotheses
 - but they criticize the Bayesian approach because it takes into account priors explicitly !

Better methods based on 'sand' than methods based on nothing!

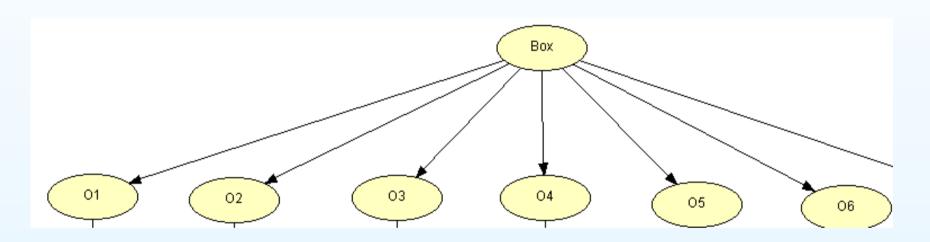
Cause-effect representation

box content \rightarrow observed color



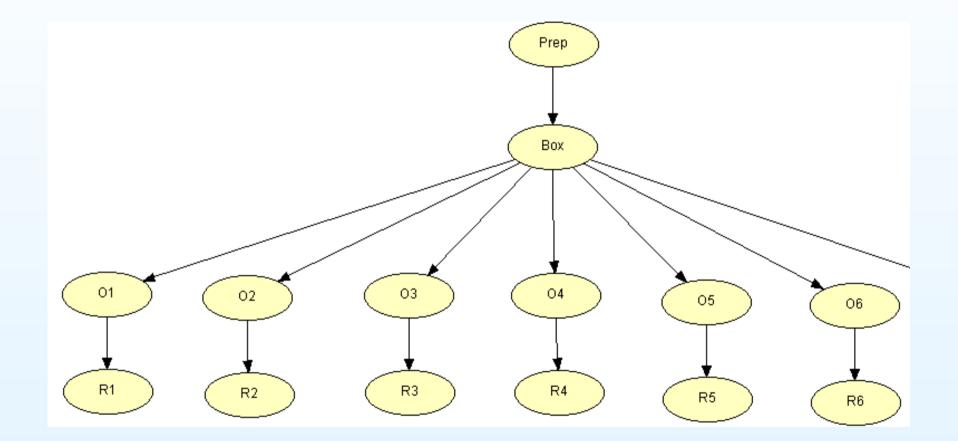
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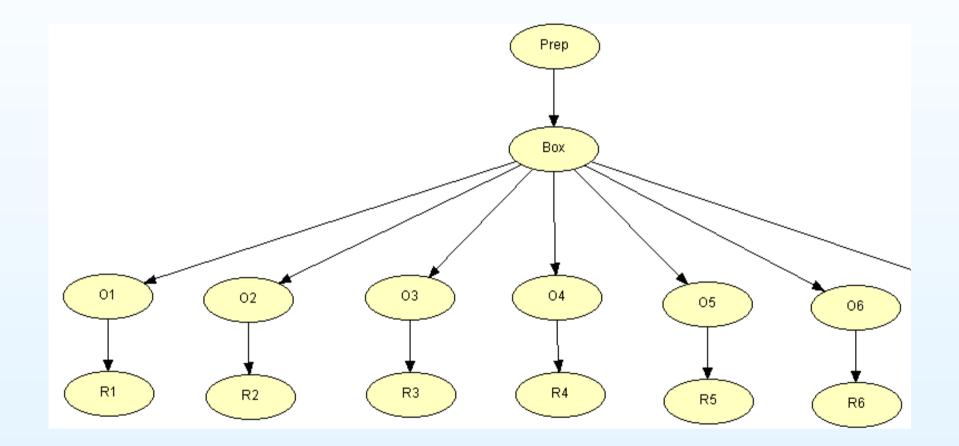


An effect might be the cause of another effect

A network of causes and effects



A network of causes and effects



and so on...

 \Rightarrow Let's play with Hugin

from 6 boxes to 1001 boxes

Overview

- example with 1001 boxes
- from uncertainty on H_j to uncertainty on p_j (proportion): $P(H_j) \leftrightarrow P(p_j)$
- Physical meaning of p_j
- Probability Vs 'Chance' ('propension')
- The discretized Bayes billard.
- The extension to continuous values of p

We are often uncertain in numbers and, consistently, we quantify of belief with probability.

Uncertain number is the more general term for random variable, though the adjective random is more committing, since it rely on the concept of randomness (see von Mises).

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- The first number rolling a die
- The temperature at the Rome airport (FCO) tomorrow at 7:00 am
- The height of the next person who enters this room

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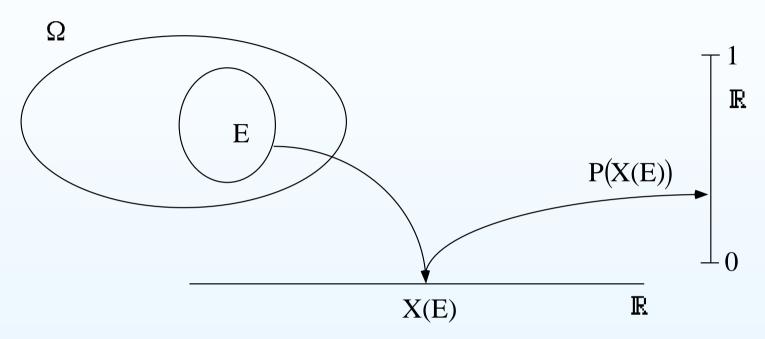
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A number respect to which we are in condition of uncertainty

- No need that the numbers can be framed in a von Mises' collective
- But it must be a well defined number (any uncertainty on its definition will increase our uncertainty about it)

From events to uncertain numbers

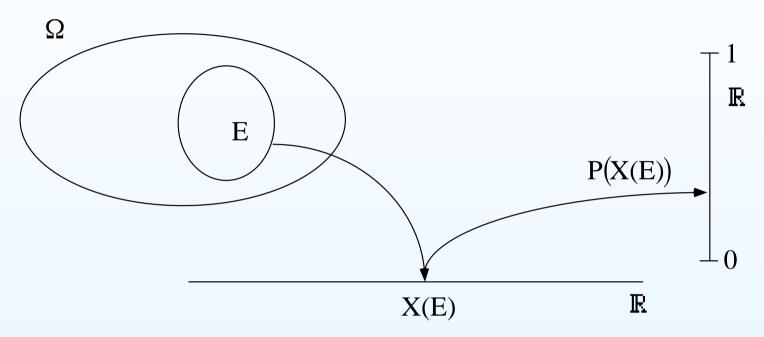


Uncertain numbers are associated to events

• Rolling one die: $X = 4 \leftrightarrow$ 'face marked with 4' (note: no intrinsic order in the numbers associated to a die)

$$\rightarrow P(X = 4) = P(\text{'face marked with 4'})$$

From events to uncertain numbers



Uncertain numbers are associated to events

 $Event \rightarrow number:$ univocal, but not bi-univocal

• Rolling two dice, with X 'sum of results'

$$\rightarrow P(X = 4) = \sum P(\text{`events giving 4'})$$

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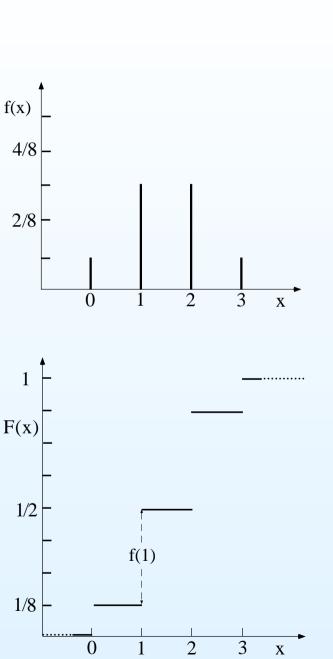
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Cumulative function (defined for all x)

$$F(x_k) \equiv P(X \le x_k) = \sum_{x_i \le x_k} f(x_i).$$

 $[F(-\infty) = 0; F(+\infty) = 1;$ $F(x_i) - F(x_{i-1}) = f(x_i);$ $\lim_{\epsilon \to 0} F(x + \epsilon) = F(x)]$



First intro to Monte Carlo

How to generate numbers in a way that their chance of occurring are proportional to f(x)?

- simple consideration based on the graphical representation of
 - $\circ f(x)$
 - $\circ F(x)$
 - extention to continuous functions
- a curious game, trowing stones...

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- The drunk man problem
 - Six keys (like rolling a die)
 - After each trial he 'loses memory'
 - We watch him and cynically bet on the attempt on which he will succeed:
 - $X = 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, \dots$

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 - $X = 1, 2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12, 13, 14, 15, \dots$
 - $\rightarrow~$ On which number would you bet?

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- \Rightarrow Beliefs are framed in a network!
 - Once we assess something, we are implicitly making an infinite number of assessments concerning logically connected events!
 - We only need to make them explicit, using logic:

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• In this case, simply chain rule:

$$P(X = 2) = P(\overline{E}_1 \cap E_2) = P(\overline{E}_1) \cdot P(E_2 | \overline{E}_1);$$

$$P(X = 3) = P(\overline{E}_1) \cdot P(\overline{E}_2 | \overline{E}_1) \cdot P(E_3 | \overline{E}_1, \overline{E}_2);$$

etc.

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[BUT sometimes the math might be hard:

→ fortunatly, nowadays most tough 'direct probability' problems can be easily solved by simulation ("Monte Carlo" methods)]

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$$f(x) = p(1-p)^{x-1}$$

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$$\dots \dots$$

$$f(x) = p(1 - p)^{x - 1}$$
Beliefs decrease geometrically

0.00

1 3 5 7 9 11 13 15 17 19 21 23 25 27 29 x

 \Rightarrow Geometric distribution [p = 1/6]



Building up $f(\boldsymbol{x})$ of the drunk man problem

p

$$f(1) = P(E_1) = p$$

$$f(2) = P(\overline{E}_1) \cdot P(E_2 | \overline{E}_1) = (1 - p) p$$

$$f(3) = P(\overline{E}_1) \cdot P(\overline{E}_2 | \overline{E}_1) \cdot P(E_3 | \overline{E}_1, \overline{E}_2) = (1 - p)^2 p$$

$$\dots \dots$$

$$f(x) = p(1 - p)^{x - 1}$$

$$p = 1/2 \rightarrow \text{tossing a coin}$$

$$[\text{Note: } f(x) = (\frac{1}{2})^x]$$

Building up f(x) of the drunk man problem

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$$\dots \dots$$

$$f(x) = p(1-p)^{x-1}$$

$$p=1/18 \rightarrow \text{a particular}$$
number at the Italian lotto
$$(p = 5/90)$$

0.01

0.00

1 3 5 7 9 11 13 15 17 19 21 23 25 27 29 х

Building up $f(\boldsymbol{x})$ of the drunk man problem

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$$\dots \dots$$

$$f(x) = p(1 - p)^{x - 1}$$
Most probable value does not depend on p .
Not a suitable indicator to state our expectation

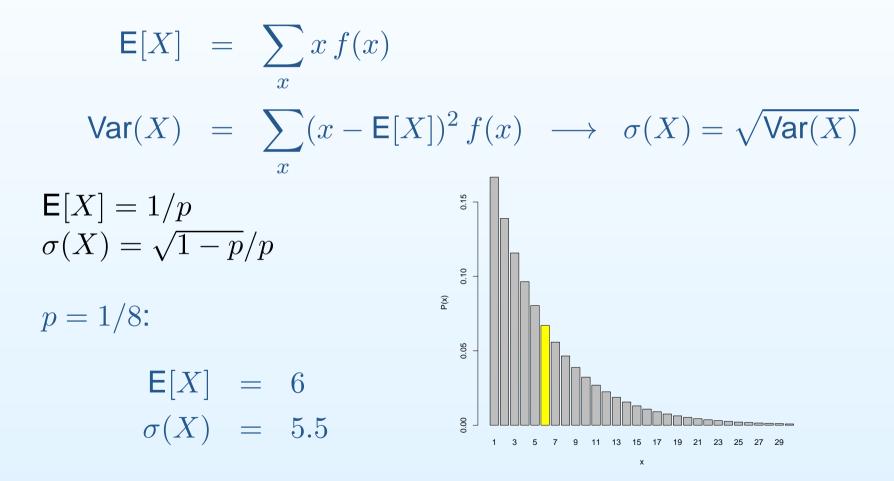
sta The same is true for the range of possibilities: $X: 1, 2, \ldots, \infty$

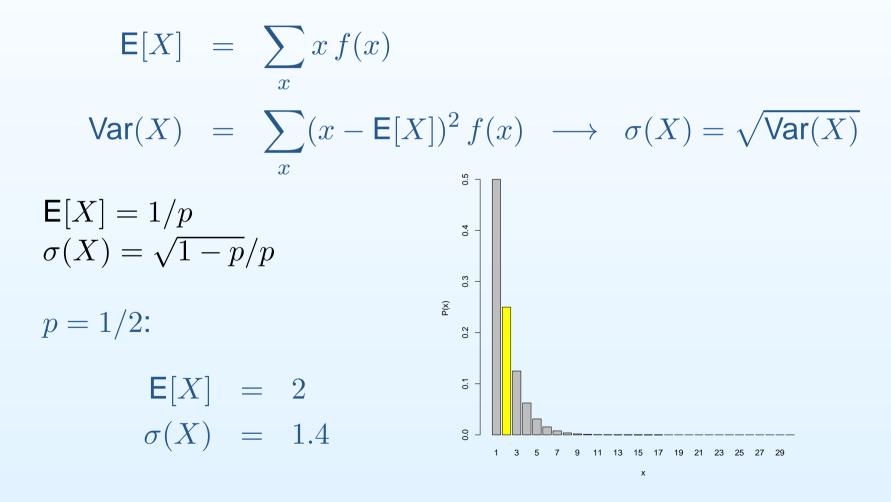
0.00 1 3 5 7 9 11 13 15 17 19 21 23 25 27 29 х

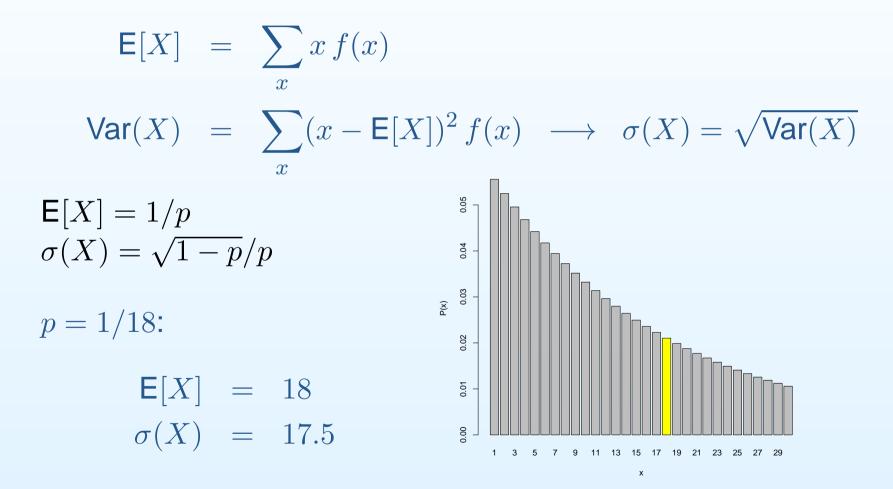
$$\mathsf{E}[X] = \sum_{x} x f(x)$$

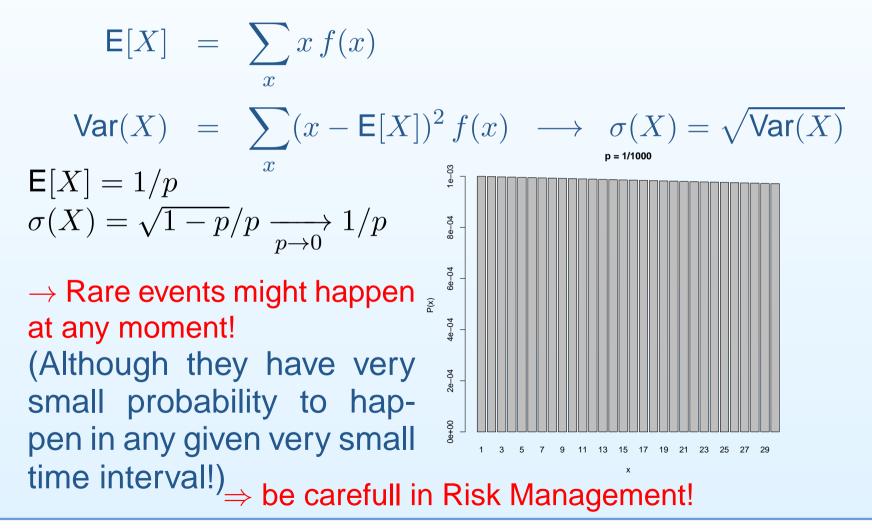
$$\mathsf{E}[X] = \sum_{x} x f(x)$$

$$\mathsf{Var}(X) = \sum_{x} (x - \mathsf{E}[X])^2 f(x) \longrightarrow \sigma(X) = \sqrt{\mathsf{Var}(X)}$$





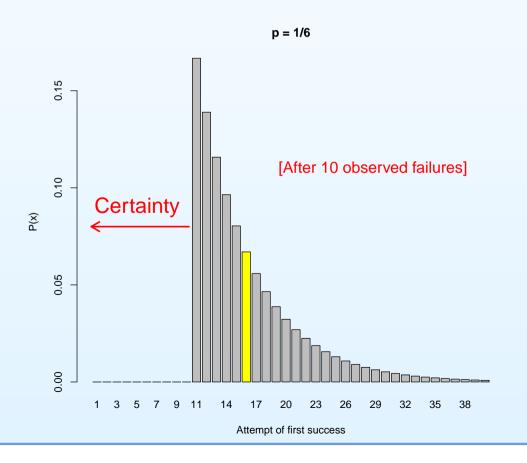




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- If $p \to 0$ in an observational interval $\Delta t \to 0$, then it makes no sense to speak about the probability in the *i*-th trial:
 - \Rightarrow It makes only sense of the probability that the 'success' occures between t_1 and t_2 ;
 - $\Rightarrow p \rightarrow \text{intensity of the Poisson process: } r = dp/dt;$
 - \Rightarrow Geometric distribution \rightarrow exponential distribution;

- If p → 0 in an observational interval Δt → 0, then it makes no sense to speak about the probability in the *i*-th trial:
 ⇒ Geometric distribution → exponential distribution;
- No memory property of geometric and exponential:



- If $p \to 0$ in an observational interval $\Delta t \to 0$, then it makes no sense to speak about the probability in the *i*-th trial:
 - \Rightarrow Geometric distribution \rightarrow exponential distribution;
- No memory property of geometric and exponential;
- Be careful: p (or r) might depend on time (think to aging effects!):
 - ⇒ geometrical/exponential model might not be any longer suitable models!

- E[X] and $\sigma(X)$ are just convenient summaries.
- In the general case they do not convey a precise confidence that X will occur in the range $E[X] \pm \sigma(X)$, though this probability is rather 'high' for typical f(x) of interest.

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When asked about the drunk man problem, most people ask they would bet on the 8-th trial, or something around it. ...and even when they are told the should bet on the first one, they reply that the first attempt has a little probability...

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- Bet on the 1-st if you win/lose if you hit/miss the number
- **BUT** sometimes wins who gets closest.
 - Bet on median if loss is linear with the error.
 - Bet on average if loss is quadratic with the error

Probability distributions Vs 'statistical' distributions

It is important to stress the difference between

- Probability distribution
 - To each possible outcome we associate how much we are confident on it:

$$x \longleftrightarrow f(x)$$

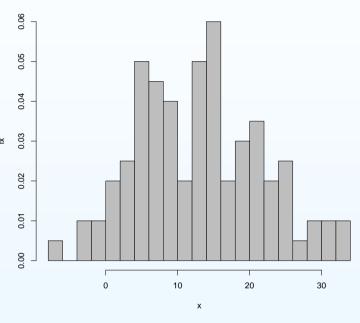
- Statistical distribution
 - To each observed outcome we associated its (relative) frequency

$$x \longleftrightarrow f_x$$

(e.g. an histogram of experimental observations) Summaries ('mean', variance, ' σ ', 'skewness', etc) have similar names and analogous definitions, but conceptual different meaning.

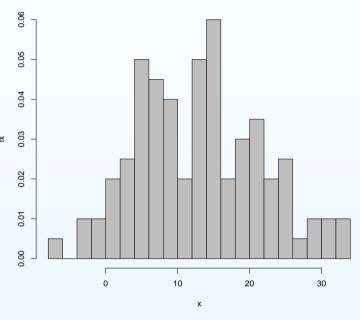
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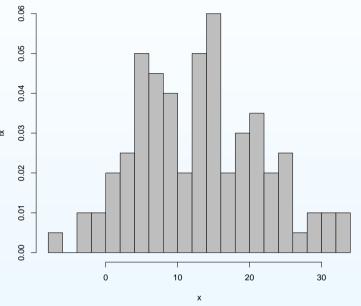


Average and variance

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$$\sigma^{2} = \sum_{x} (x - \overline{x})^{2} f_{x}$$

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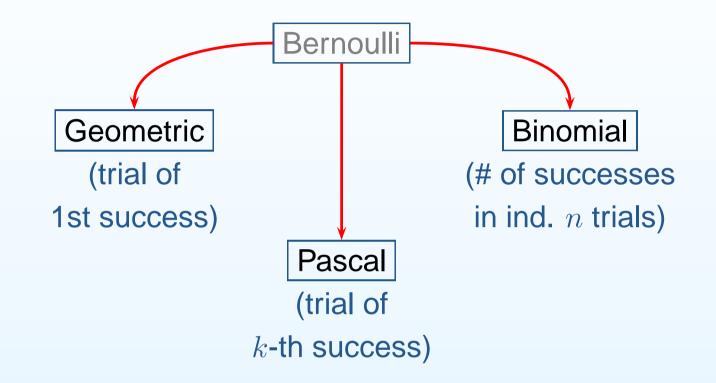


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→ Just a rough empirical description of the shape \Rightarrow center of mass and momentum of inertia! (Famous 'n/(n-1)' correction: interference descriptive \leftrightarrow inferential statistics.)

Distributions derived from the Bernoulli process



(Binomial well known. We shall not use the Pascal)

Poisson distribution

One of the best known distributions by physicist.

For a while, just take the mathematical approach to the Poisson distribution:

$$f(x \mid \mathcal{P}_{\lambda}) = \frac{\lambda^{x}}{x!} e^{-\lambda} \qquad \begin{cases} 0 < \lambda < \infty \\ x = 0, 1, \dots, \infty \end{cases}$$

Reminding also the well known property

$$\mathcal{B}_{n,p} \xrightarrow{n \to \infty} P_{\lambda} .$$

$$n \to \infty$$

$$p \to 0$$

$$(n \, p = \lambda)$$

Poisson process



Let us consider some phenomena that might happen at a give instant, such that

• Probability of 1 count in ΔT is proportional to ΔT , with ΔT 'small'.

$$p = P($$
"1 count in $\Delta T'') = r \Delta T$

where r is the intensity of the process'

- $P(\geq 2 \text{ counts}) \ll P(1 \text{ count})$ (OK if ΔT is small enough)
- What happens in one interval does not depend on other intervals (if disjoints)

Let us divide a finite interval *T* in *n* small intervals, i.e. $T = n \Delta T$, and $\Delta T = T/n$.

Poisson process \rightarrow Poisson distribution



Considering the possible occurrence of a count in each small interval ΔT an independent Bernoulli trial, of probability

$$p = r\,\Delta T = r\,T/n$$

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But $n \to \infty$ and $p \to 0 \Rightarrow \mathcal{B}_{n,p} \to \mathcal{P}_{\lambda}$ where $\lambda = n p = r T$

 $\Rightarrow \lambda$ depends only on the intensity of the process and on the finite time of observation.

Poisson process \rightarrow waiting time



Another interesting problem: how long do we have to wait for the first count? (Starting from any arbitrary time)

Problem analogous to the Geometric, but now it makes no sense to ask at which small interval the counts will occur!

Poisson process \rightarrow waiting time



Another interesting problem: how long do we have to wait for the first count? (Starting from any arbitrary time)

Problem analogous to the Geometric, but now it makes no sense to ask at which small interval the counts will occur!

Let us restart from the Geometric and calculate P(X > x):

$$P(X > x) = \sum_{i > x} f(i | \mathcal{G}_p) = (1 - p)^x$$

(The count will not occur in the first x trials).

In the domain of time, using p = r t/n and then making the limit:

$$P(T > t) = (1 - p)^n = (1 - r t/n)^n \xrightarrow[n \to \infty]{} e^{-rt}$$

Poisson process \rightarrow Exponential distribution

Knowing P(T > t) we get easily the cumulative F(t):

$$F(t) = P(T \le t) = 1 - P(T > t) = 1 - e^{-rt}$$

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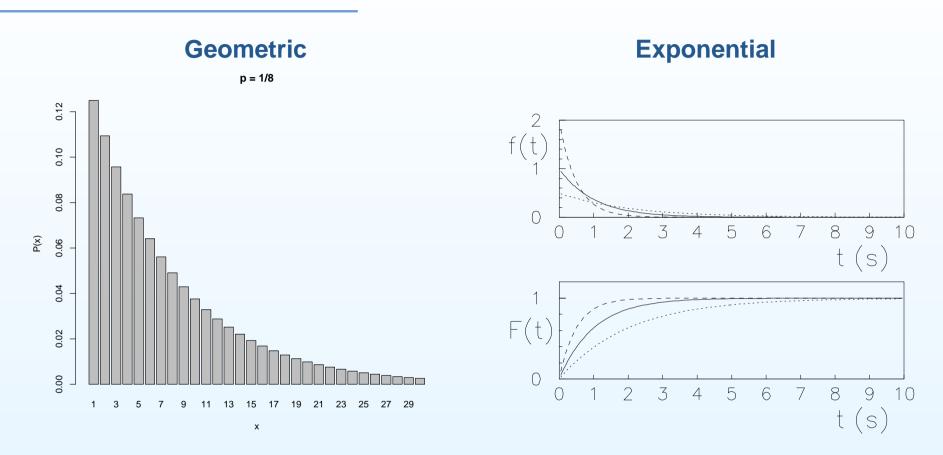
→ This leads us to define a probability density function (pdf) for continuous variables:

$$f(t) = \frac{d F(x)}{d t} \,.$$

- In this case $f(t) = r e^{-rt} = \frac{1}{\tau} e^{-t/\tau}$
- \rightarrow Exponential distribution ($\tau = 1/r$): $E[T] = \sigma(T) = \tau$.

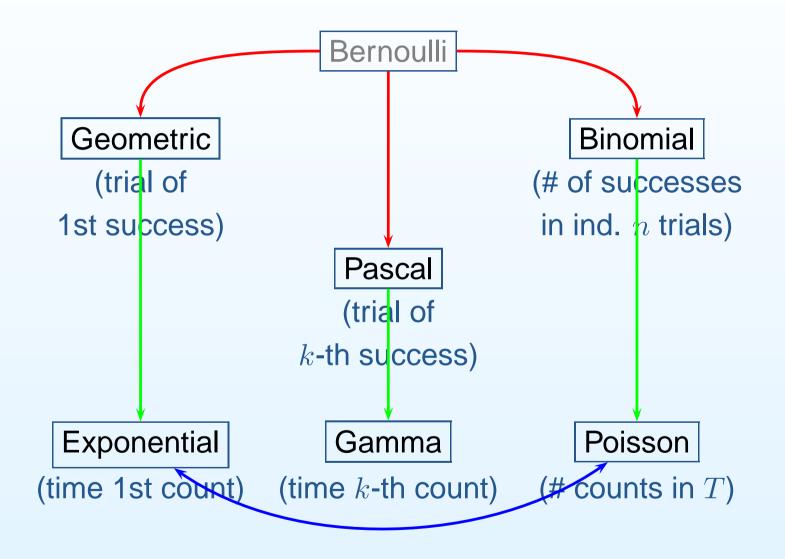
 $(\Rightarrow$ Properties of pdf assumed to be well known.)

Geometric \leftrightarrow Exponential

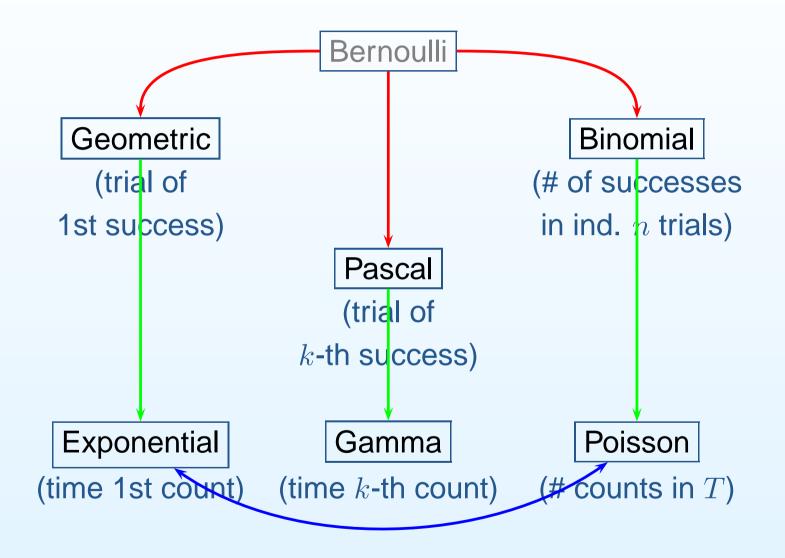


Exponential is just the limit to the continuum of the Geometric. 'No memory' property for both: Assuming a success (or a count) has not happened until a certain trial (or time), the distributions restart from there. No need to know the instant of particle creation to measure 'life time' (\rightarrow the "10³³ year old" proton!).

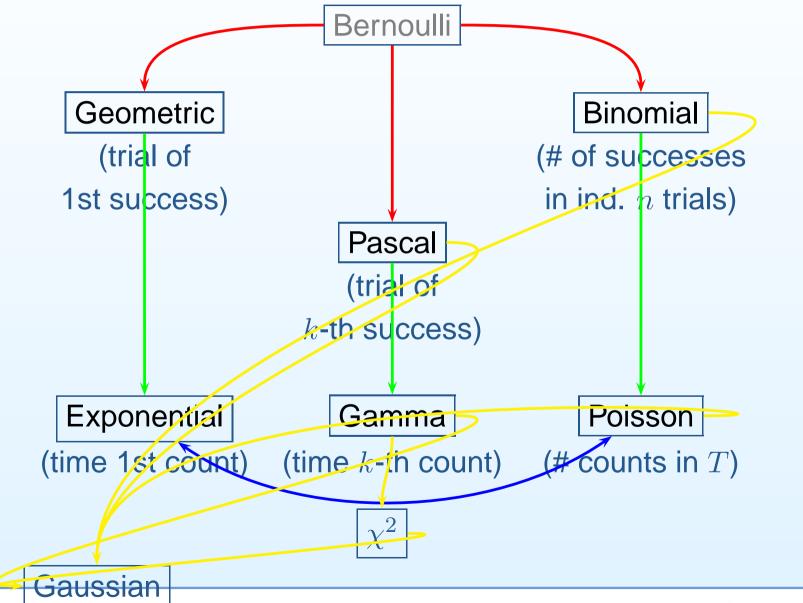
Distributions derived from the Bernoulli process



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Distributions derived from the Bernoulli process



Note

Though we could not go through all technical details, it is important to remark that all these distributions are obtained assuming that each 'act of observation', that can be asymptotically associated to a single point, is an independent Bernoulli trial of constant probability p (that might tend to zero). Important properties of probability distributions

 $E(\cdot)$ is a linear operator:

$$\mathsf{E}(aX+b) = a\,\mathsf{E}(X) + b\,.$$

Transformation properties of variance and standard deviation:

$$Var(aX + b) = a^2 Var(X),$$

$$\sigma(aX + b) = |a| \sigma(X).$$

Obviously, I have to assume that most of the basic formalism is well known, e.g. that $P(a \le X \le b) = \int_a^b f(x) dx$, etc.

From probability to future frequencies

Let us think to n independent Bernoulli trials that <u>have to be made</u>.

Number of successes $X \sim \mathcal{B}_{n,p}$, with p.

We might be interested to the relative frequency of successes, i.e. $f_n = X/n$: $f_n = 0, 1/n, 2/n, ..., 1$

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What do we expect for f_n ? $f(f_n)$ can be obtained from f(x).

$$\mathsf{E}(f_n) \equiv \frac{1}{n} \mathsf{E}(X \mid \mathcal{B}_{n,p}) = \frac{n p}{n} = p$$

$$\sigma(f_n) \equiv \frac{1}{n} \sigma(X \mid \mathcal{B}_{n,p}) = \frac{\sqrt{p (1-p)}}{\sqrt{n}} \xrightarrow[n \to \infty]{} 0$$

We expect p, with uncertainty that decreases with \sqrt{n} : \rightarrow *Bernoulli's theorem*, the most known, misunderstood and misused probability theory theorem.

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In particular, it justifies the increased probability of neither 'late numbers' at lotto, nor frequency based definition of probability (Circular: cannot define probability from probability theorem!)

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 - $^{\circ}$ "Probabilità come propensione" \approx OK

 $\rightarrow P(E | \mathsf{prop} = p) = p$

 $\rightarrow f_n$ dati eventi analoghi in cui crediamo che la propensione sia la stessa, vale il Th. di Bernoulli.

Probabilità <u>come</u> limite della frequenza: NO

 Non si dimentichi che il teorema di Bernoulli ... è un teorema, basato sulle regole di base dalla probabilità e su tutte le proprietà che ne derivano. Quindi non può definire il concetto di probabilità.

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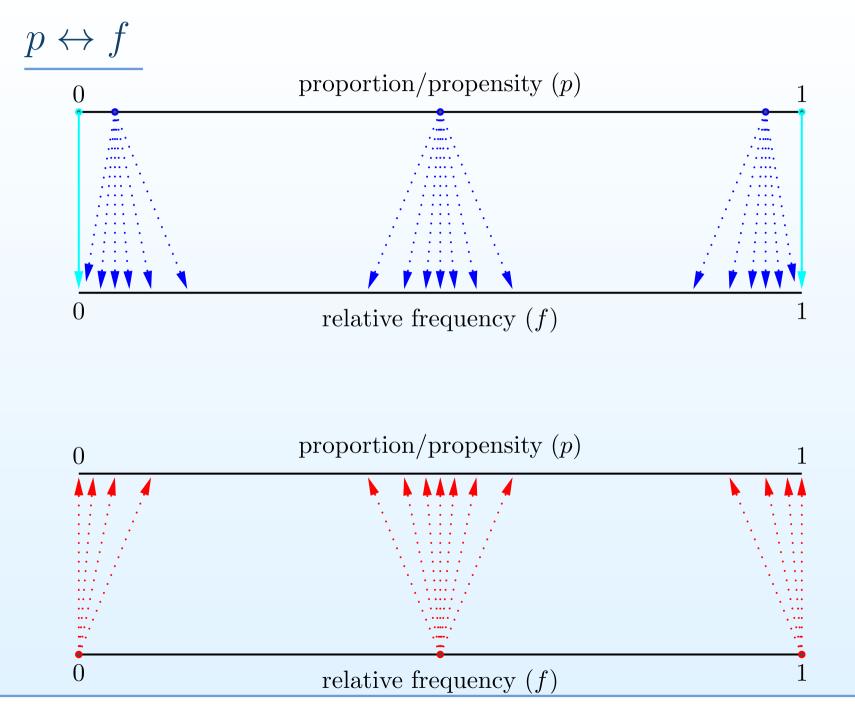
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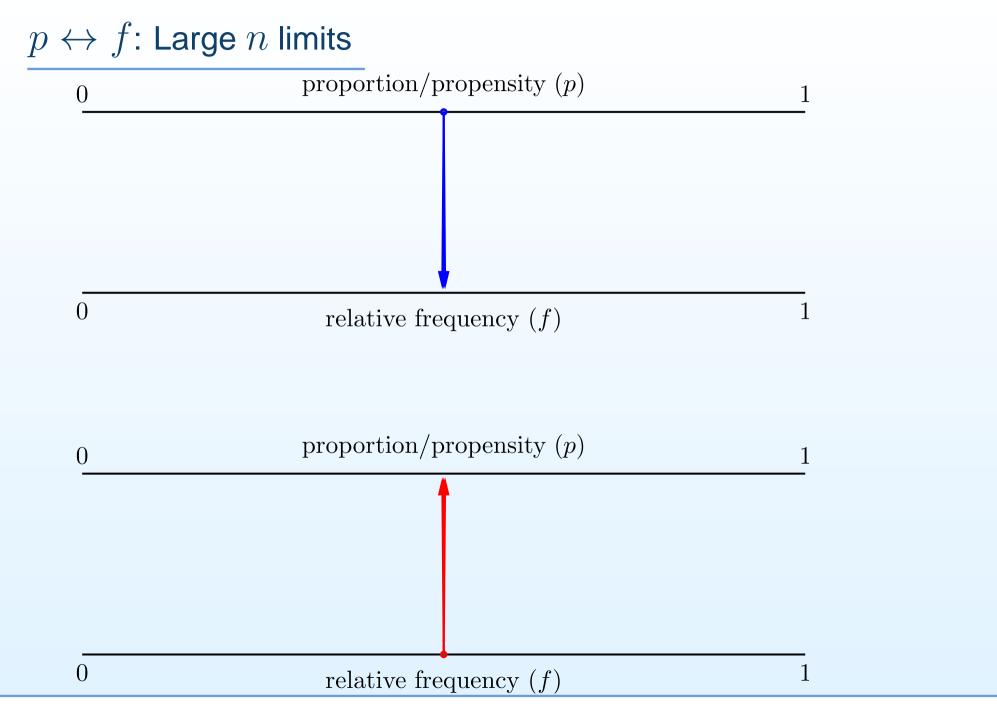
 Si noti inoltre che la condizione di p costante implica che essa sia prefissata a priori e che anche le valutazioni sui possibili esiti di f_n siano fatte prima di iniziare le prove (o in condizione di incertezza sul loro esito). Probabilità Vs frequenza relativa

Concetti da tenere ben distinti anche se esiste un collegamento fra di loro:

 $p \rightarrow f_n$: Teorema di Bernoulli

 $f_n \rightarrow p$: Teorema di Bayes sotto precise condizioni





Propagation of uncertainties

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 \Rightarrow Therefore, the famous problem of propagation of uncertainty is straightforward in a probabilistic approach: just use probability theory.

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The general problem:

$$f(x_1, x_2, \ldots, x_n) \xrightarrow{Y_j = Y_j(X_1, X_2, \ldots, X_n)} f(y_1, y_2, \ldots, y_m).$$

This calculation can be quite challenging, but it can be easily performed by Monte Carlo techniques.

General solution for discrete variables

Y = Y(X), where Y() stands for the mathematical function relating X and Y.

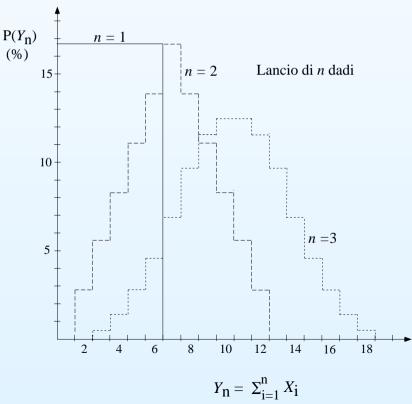
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Probability distributions of the sums of the results from n dice.



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The extension to many variables is straightforward: for ex., given two *input* quantities X_1 and X_2 , with their probability function $f(x_1, x_2)$, and two *output* quantities Y_1 and Y_2 :

$$f(y_1, y_2) = \sum_{\substack{x_1, x_2 \\ Y_1(x_1, x_2) = y_1 \\ Y_2(x_1, x_2) = y_2}} f(x_1, x_2)$$

(For each point $\{y_1, y_2\}$ sum up the probability of all points in the $\{X_1, X_2\}$ space that satisfy the constrain.)

Just extend to the continuum the previous formula:

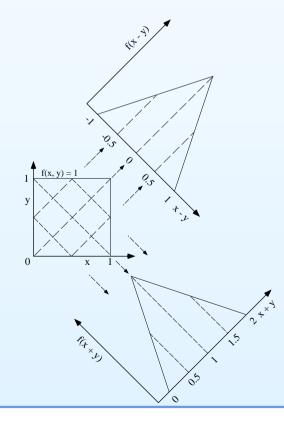
- replace sums by integrals
- replace constrains by suitable Dirac $\delta()$:

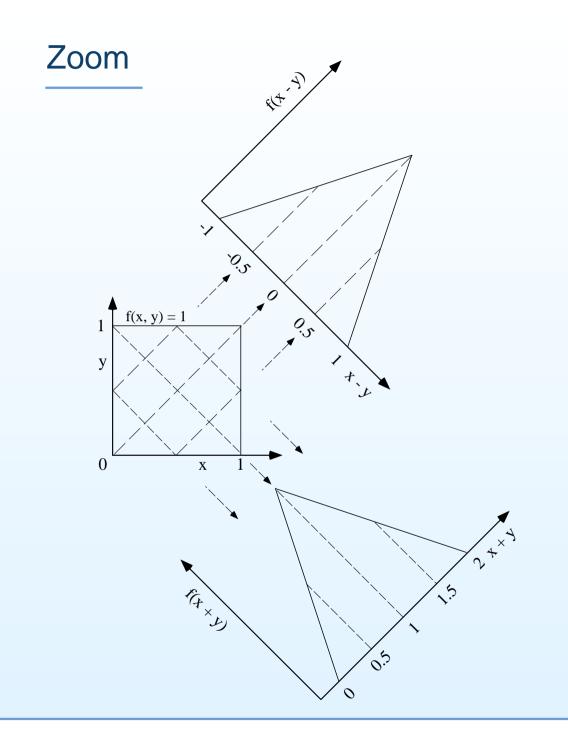
 $f(y_1, y_2) = \int \delta(y_1 - Y_1(x_1, x_2)) \,\delta(y_2 - Y_2(x_1, y_2)) \,f(x_1, x_2) \,\mathrm{d}x_1 \mathrm{d}x_2 \,.$

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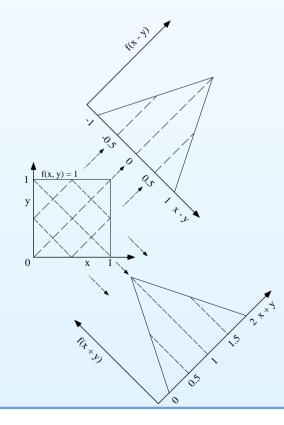




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- replace constrains by suitable Dirac $\delta()$:

 $f(y_1, y_2) = \int \delta(y_1 - Y_1(x_1, x_2)) \,\delta(y_2 - Y_2(x_1, y_2)) \,f(x_1, x_2) \,\mathrm{d}x_1 \mathrm{d}x_2 \,.$

$$\mathsf{E}(Y) = \mathsf{E}(X_1) + \mathsf{E}(X_2)$$

$$\sigma^2(Y) = \sigma^2(X_1) + \sigma^2(X_2)$$

 $mode(Y) \leftrightarrow mode(X_i)$ $median(Y) \leftrightarrow median(X_i)$

?

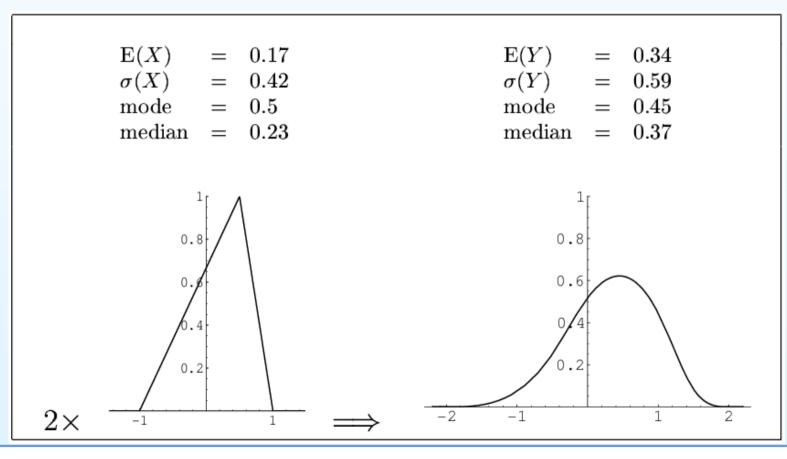
$\mathrm{E}(X)$	=	0.17
$\sigma(X)$	=	0.42
mode	=	0.5
median	=	0.23

E(Y) = 0.34 $\sigma(Y) = 0.59$ mode = 0.45 median = 0.37

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And remember that standard methods (χ^2 or ML fits) provide something equivalent to 'most probable values', not to E()!

(As we shall see.)



Notes

The following slides should be reached by hyper-links, clicking on words with the symbol †

If I eat a chicken and you eat no chicken...

... for the statistics each of us eats 1/2 chicken. For the pleasure of Italian readers, this is how Trilussa put is:

La statistica

Sai ched'è la statistica? È 'na cosa che serve pe' fa' un conto in generale de la gente che nasce, che sta male, che more, che va in carcere e che sposa.

Ma pe' me la statistica curiosa è dove c'entra la percentuale, pe' via che, lì, la media è sempre eguale puro co' la persona bisognosa.





Me spiego, da li conti che se fanno seconno le statistiche d'adesso risurta che te tocca un pollo all'anno:

e, se nun entra ne le spese tue, t'entra ne la statistica lo stesso perché c'è un antro che se ne magna due.



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 - $^{\circ}$ And even statistics experts, when they have to transmit to the rest of the community the meaning of what they do, they have hard time in doing it \longrightarrow Slide

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 - At least, clear after 2000 CERN CLW → Slide (But I am afraid if I would redo the survey now, I would get similar answers...)

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Notwithstanding the fact that there is been a lot of activity in the past years by several physicists, convinced that the idea is basically good, but one only needs 'a better prescription'.

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If the method guarantees the claimed coverage, who refunds us if it does not work?

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- After six years the production of 90-95% C.L. bounds has continued steadly, and in many cases the so called 'unified approach' has been used, but still coverage does not do its job.
- What will be the next excuse?
- \Rightarrow I do not know what the so-called 'flip-plopping' is,

but we can honestly acknowledge the flop of that reasoning.

