$$f(\lambda \,|\, x, \mathcal{P}) = \frac{\frac{\lambda^{x} e^{-\lambda}}{x!} f_{\circ}(\lambda)}{\int_{0}^{\infty} \frac{\lambda^{x} e^{-\lambda}}{x!} f_{\circ}(\lambda) d\lambda}.$$



$$f(\lambda \,|\, x, \mathcal{P}) = \frac{\frac{\lambda^{x} e^{-\lambda}}{x!} f_{o}(\lambda)}{\int_{0}^{\infty} \frac{\lambda^{x} e^{-\lambda}}{x!} f_{o}(\lambda) d\lambda}.$$

Assuming $f_{\circ}(\lambda)$ constant up to a certain $\lambda_{max} \gg x$ and making the integral by parts we obtain

$$f(\lambda \mid x, \mathcal{P}) = \frac{\lambda^{x} e^{-\lambda}}{x!}$$
$$F(\lambda \mid x, \mathcal{P}) = 1 - e^{-\lambda} \left(\sum_{n=0}^{x} \frac{\lambda^{n}}{n!} \right)$$



$$f(\lambda \,|\, x, \mathcal{P}) = \frac{\frac{\lambda^{x} e^{-\lambda}}{x!} f_{\circ}(\lambda)}{\int_{0}^{\infty} \frac{\lambda^{x} e^{-\lambda}}{x!} f_{\circ}(\lambda) d\lambda}.$$

Assuming $f_{\circ}(\lambda)$ constant up to a certain $\lambda_{max} \gg x$ and making the integral by parts we obtain

$$f(\lambda | x, \mathcal{P}) = \frac{\lambda^{x} e^{-\lambda}}{x!}$$
$$F(\lambda | x, \mathcal{P}) = 1 - e^{-\lambda} \left(\sum_{n=0}^{x} \frac{\lambda^{n}}{n!} \right)$$

Summaries

$$E(\lambda) = x + 1,$$

$$Var(\lambda) = x + 1,$$

$$\lambda_m = x$$

© GdA, RM25-13 17/02/25 10/37

Some examples of $f(\lambda)$



For 'large' $x f(\lambda | x)$ becomes Gaussian with expected value x and standard deviation \sqrt{x} .

The difference between the most probable λ and its expected value for small x is due to the asymmetry of $f(\lambda)$.

Conjugate prior

$f(\lambda | x) \propto \lambda^{x} e^{-\lambda} \cdot f_{o}(\lambda)$

© GdA, RM25-13 17/02/25 14/37

Conjugate prior

$$egin{aligned} f(\lambda \,|\, x) &\propto &\lambda^x \, e^{-\lambda} \cdot f_{\circ}(\lambda) \ &\propto &\lambda^x \, e^{-\lambda} \cdot \lambda^a \, e^{-b\,\lambda} \end{aligned}$$

C GdA, RM25-13 17/02/25 14/37

Conjugate prior

$$egin{aligned} f(\lambda \,|\, x) &\propto &\lambda^x \, e^{-\lambda} \cdot f_{
m o}(\lambda) \ &\propto &\lambda^x \, e^{-\lambda} \cdot \lambda^a \, e^{-b\,\lambda} \ &\propto &\lambda^{x+a} \, e^{-(1+b)\,\lambda} \end{aligned}$$

Does such a probability function 'exist'?

© GdA, RM25-13 17/02/25 14/37

Conjugate prior

$$egin{aligned} f(\lambda \,|\, x) &\propto &\lambda^x \, e^{-\lambda} \cdot f_{
m o}(\lambda) \ &\propto &\lambda^x \, e^{-\lambda} \cdot \lambda^a \, e^{-b\,\lambda} \ &\propto &\lambda^{x+a} \, e^{-(1+b)\,\lambda} \end{aligned}$$

Does such a probability function 'exist'?

 \Rightarrow Gamma distribution



$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).



$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

c is called *shape* parameter, while 1/r is the *scale* parameter.



$$X \sim \text{Gamma}(c, r):$$

$$f(x | \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

- c is called *shape* parameter, while 1/r is the *scale* parameter.
 - If c is integer, the distribution is also known as Erlang, describing the time to wait before observing the c-th event in a Poisson process of intensity ('rate') r.

$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

c is called *shape* parameter, while 1/r is the *scale* parameter.

- If c is integer, the distribution is also known as Erlang, describing the time to wait before observing the c-th event in a Poisson process of intensity ('rate') r.
- For c = 1 the Gamma distribution recovers the exponential.

$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

c is called *shape* parameter, while 1/r is the *scale* parameter.

- If c is integer, the distribution is also known as Erlang, describing the time to wait before observing the c-th event in a Poisson process of intensity ('rate') r.
- For c = 1 the Gamma distribution recovers the exponential.
- Finally, the χ^2 distribution is just a particular Gamma:

$$f(x | \chi_{\nu}^2) = f(x | \text{Gamma}(\nu/2, 1/2))$$

$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

c is called *shape* parameter, while 1/r is the *scale* parameter.

- If c is integer, the distribution is also known as Erlang, describing the time to wait before observing the c-th event in a Poisson process of intensity ('rate') r.
- For c = 1 the Gamma distribution recovers the exponential.
- Finally, the χ^2 distribution is just a particular Gamma:

$$f(x | \chi_{\nu}^2) = f(x | \text{Gamma}(\nu/2, 1/2))$$

► The Gamma is a key distribution!

$$X \sim \text{Gamma}(c, r):$$

$$f(x \mid \text{Gamma}(c, r)) = \frac{r^{c}}{\Gamma(c)} x^{c-1} e^{-rx} \qquad \left\{ \begin{array}{l} r, c > 0 \\ x \ge 0 \end{array} \right.,$$

where

$$\Gamma(c) = \int_0^\infty x^{c-1} e^{-x} \mathrm{d}x$$

(for *n* integer, $\Gamma(n+1) = n!$).

c is called *shape* parameter, while 1/r is the *scale* parameter.

- If c is integer, the distribution is also known as Erlang, describing the time to wait before observing the c-th event in a Poisson process of intensity ('rate') r.
- For c = 1 the Gamma distribution recovers the exponential.
- Finally, the χ^2 distribution is just a particular Gamma:

$$f(x | \chi_{\nu}^2) = f(x | \text{Gamma}(\nu/2, 1/2))$$

The Gamma is a key distribution! The Erlang distribution is important to get a physical intuition of the properties of Gamma and then of the χ²!

Some examples



r: rate (if the variable is a time, then *r* is Poisson rate).

Some examples



r: rate (rate increases \rightarrow distributions squized)

Some examples



r: rate (rate increases \rightarrow distributions squized)

Gamma (and χ^2) distribution $_{\rm Summaries}$

$$E(X) = \frac{c}{r}$$

$$Var(X) = \frac{c}{r^2} = \frac{E(X)}{r}$$

$$mode(X) = \begin{cases} 0 & \text{if } c \le 1 \\ \frac{c-1}{r} & \text{if } c > 1 \end{cases}$$

C GdA, RM25-13 17/02/25 19/37

Gamma (and χ^2) distribution Summaries

$$E(X) = \frac{c}{r}$$

$$Var(X) = \frac{c}{r^2} = \frac{E(X)}{r}$$

$$mode(X) = \begin{cases} 0 & \text{if } c \le 1 \\ \frac{c-1}{r} & \text{if } c > 1 \end{cases}$$

Therefore, for the χ^2 (ightarrow c =
u/2, r = 1/2)

$$E(\chi^2) = \nu$$

$$Var(\chi^2) = 2\nu$$

$$mode(\chi^2) = \begin{cases} 0 & \text{if } \nu \leq 2\\ \nu - 2 & \text{if } \nu > 2 \end{cases}$$

© GdA, RM25-13 17/02/25 19/37

Distributions derived from the Bernoulli process



C GdA, RM25-13 17/02/25 20/37

Distributions derived from the Bernoulli process



C GdA, RM25-13 17/02/25 20/37

Distributions derived from the Bernoulli process



C GdA, RM25-13 17/02/25 20/37

Use of gamma conjugate prior

$f(\lambda | x, \text{Gamma}(c_i, r_i)) \propto [\lambda^x e^{-\lambda}] \times [\lambda^{c_i-1} e^{-r_i \lambda}]$



Use of gamma conjugate prior

$$egin{aligned} f(\lambda \,|\, x, \operatorname{Gamma}(c_i, r_i)) & \propto & \left[\lambda^x e^{-\lambda}
ight] imes \left[\lambda^{c_i-1} e^{-r_i\,\lambda}
ight] \ & \propto & \lambda^{x+c_i-1} e^{-(r_i+1)\,\lambda}\,, \end{aligned}$$

where c_i and r_i are the initial parameters of the gamma distribution.



Use of gamma conjugate prior

$$egin{aligned} f(\lambda \,|\, x, \operatorname{Gamma}(c_i, r_i)) & \propto & \left[\lambda^x e^{-\lambda}
ight] imes \left[\lambda^{c_i-1} e^{-r_i\,\lambda}
ight] \ & \propto & \lambda^{x+c_i-1} e^{-(r_i+1)\,\lambda}\,, \end{aligned}$$

where c_i and r_i are the initial parameters of the gamma distribution.

Updating rule

 $c_f = c_i + x$ $r_f = r_i + 1$

© GdA, RM25-13 17/02/25 21/37

Use of gamma conjugate prior

$$egin{aligned} f(\lambda \,|\, x, \operatorname{Gamma}(c_i, r_i)) & \propto & \left[\lambda^x e^{-\lambda}
ight] imes \left[\lambda^{c_i-1} e^{-r_i\,\lambda}
ight] \ & \propto & \lambda^{x+c_i-1} e^{-(r_i+1)\,\lambda}\,, \end{aligned}$$

where c_i and r_i are the initial parameters of the gamma distribution.

Updating rule

$$c_f = c_i + x$$

 $r_f = r_i + 1$

► A "flat conjugate" prior (not just academic!):



Use of gamma conjugate prior

$$egin{aligned} f(\lambda \,|\, x, \operatorname{Gamma}(c_i, r_i)) & \propto & \left[\lambda^x e^{-\lambda}
ight] imes \left[\lambda^{c_i-1} e^{-r_i\,\lambda}
ight] \ & \propto & \lambda^{x+c_i-1} e^{-(r_i+1)\,\lambda}\,, \end{aligned}$$

where c_i and r_i are the initial parameters of the gamma distribution.

Updating rule

$$c_f = c_i + x$$

 $r_f = r_i + 1$

A "flat conjugate" prior (not just academic!): \rightarrow exponential with very large τ (or vanishing r)



Use of gamma conjugate prior

$$egin{aligned} f(\lambda \,|\, x, \operatorname{Gamma}(c_i, r_i)) & \propto & \left[\lambda^x e^{-\lambda}
ight] imes \left[\lambda^{c_i-1} e^{-r_i\,\lambda}
ight] \ & \propto & \lambda^{x+c_i-1} e^{-(r_i+1)\,\lambda}\,, \end{aligned}$$

where c_i and r_i are the initial parameters of the gamma distribution.



$$c_f = c_i + x$$

 $r_f = r_i + 1$

 A "flat conjugate" prior (not just academic!): → exponential with very large τ (or vanishing r)
 c = 1, r → 0

$$f(\lambda | x, \text{Gamma}(c_i = 1, r_i \rightarrow 0)) \propto \lambda^x e^{-\lambda}$$