Improved (iterative) Bayesian unfolding

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Outline

- Learning from data the probabilistic way
 - \circ Causes \longleftrightarrow Effects

"The essential problem of the experimental method" (Poincaré).

- Graphical representation of probabilistic links
- Learning about causes from their effects
- Parametric inference Vs unfolding
- From principles to real life... [the iteration 'dirty trick']
- The old code and its weak point
- Improvements:
 - use (conjugate) pdf's insteads of just 'estimates'
 - uncertainty evaluated by general rules of probability (instead of 'error propagation' formulae)
 - \Rightarrow integrals over the weighted possibilities \rightarrow MC
- Some examples on toy models

Learning from experience and source of uncertainty



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Causes \rightarrow effects

The same *apparent* cause might produce several, different effects



Given an observed effect, we are not sure about the exact cause that has produced it.

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$$\mathbf{E_2} \Rightarrow \{C_1, C_2, C_3\}?$$

The essential problem of the experimental method

"Now, these problems are classified as *probability of causes*, and are most interesting of all their scientific applications. I play at *écarté* with a gentleman whom I know to be perfectly honest. What is the chance that he turns up the king? It is 1/8. This is a problem of the *probability of effects*.

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"Now, these problems are classified as *probability of causes*, and are most interesting of all their scientific applications. I play at *écarté* with a gentleman whom I know to be perfectly honest. What is the chance that he turns up the king? It is 1/8. This is a problem of the *probability of effects*.

I play with a gentleman whom I do not know. He has dealt ten times, and he has turned the king up six times. What is the chance that he is a sharper? This is a problem in the probability of causes. It may be said that it is the essential problem of the experimental method."

(H. Poincaré – Science and Hypothesis)

Uncertainty and probability

We, as physicists, consider absolutely natural and meaningful statements of the following kind

- $P(-10 < \epsilon'/\epsilon \times 10^4 < 50) >> P(\epsilon'/\epsilon \times 10^4 > 100)$
- $P(170 \le m_{top}/\text{GeV} \le 180) \approx 70\%$
- $P(M_H < 200 \,\text{GeV}) > P(M_H > 200 \,\text{GeV})$

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I stick to common sense (and physicists common sense) and assume that probabilities of causes, probabilities of of hypotheses, probabilities of the numerical values of physics quantities, etc. are sensible concepts that match the mind categories of human beings

(see D. Hume, C. Darwin + modern researches)

Our original problem:



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Our conditional view of probabilistic causation

$$P(E_i \mid C_j)$$

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The fourth basic rule of probability:

 $P(C_j, E_i) = P(E_i | C_j) P(C_j) = P(C_j | E_i) P(E_i)$

Let us take basic rule 4, written in terms of hypotheses H_j and effects E_i , and rewrite it this way:

$$\frac{P(H_j \mid E_i)}{P(H_j)} = \frac{P(E_i \mid H_j)}{P(E_i)}$$

"The condition on E_i changes in percentage the probability of H_j as the probability of E_i is changed in percentage by the condition H_j ."

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Got 'after'

Calculated 'before'

(where 'before' and 'after' refer to the knowledge that E_i is true.)

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 \Rightarrow Bayes theorem

A different way to view fit issues



- Determistic link μ_x 's to μ_y 's
- Probabilistic links $\mu_x \to x$, $\mu_y \to y$
- $\Rightarrow \text{ aim of fit: } \{ \boldsymbol{x}, \boldsymbol{y} \} \rightarrow \boldsymbol{\theta} \Rightarrow f(\boldsymbol{\theta} \,|\, \{ \boldsymbol{x}, \boldsymbol{y} \})$

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probabilistic parametric inference \Rightarrow it relies on the kind of functions parametrized by θ $\mu_y = \mu_y(\mu_x; \theta)$

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BUT sometimes we wish to interpret the data as little as possible

⇒ just public 'something equivalent' to an experimental distribution, with the bin contents fluctuating according to an underlying multinomial distribution, but having possibly got rid of physical and instrumental distortions, as well as of background.

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- \Rightarrow Unfolding (deconvolution)

Why unfolding?

Te idea is to provide something similar to an experimental spectrum, with a minimal interpretation by the experimentalist, a part from correcting from distortions due to physics and detector effects (including background).

(The alternative would be to give a parametrized description of the true spectrum – a fit)

Invert smearing matrix?

Invert smearing matrix? In general is a bad idea: not a rotational problem but an inferential problem!

Imagine
$$S = \begin{pmatrix} 0.8 & 0.2 \\ 0.2 & 0.8 \end{pmatrix}$$
: $\rightarrow U = S^{-1} = \begin{pmatrix} 1.33 & -0.33 \\ -0.33 & 1.33 \end{pmatrix}$
Let the true be $s_t = \begin{pmatrix} 10 \\ 0 \end{pmatrix}$: $\rightarrow s_m = S \cdot s_t = \begin{pmatrix} 8 \\ 2 \end{pmatrix}$;
If we measure $s_m = \begin{pmatrix} 8 \\ 2 \end{pmatrix} \rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 10 \\ 0 \end{pmatrix} \checkmark$

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Indeed, matrix inversion is recognized to producing 'crazy spectra' and even negative values (unless such large numbers in bins such fluctuations around expectations are negligeable)



(T: 'trash')



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- x_C : true spectrum (nr of events in cause bins)
- x_E : observed spectrum (nr of events in effect bins)



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- x_C : true spectrum (nr of events in cause bins)
- \boldsymbol{x}_E : observed spectrum (nr of events in effect bins)

Our aim:

- not to find the true spectrum
- but, more modestly, <u>rank in beliefs</u> all possible spectra that might have caused the observed one: $\Rightarrow P(\mathbf{x}_C | \mathbf{x}_E, I)$



• $P(\mathbf{x}_C | \mathbf{x}_E, I)$ depends on the knowledge of smearing matrix Λ , with $\lambda_{ji} \equiv P(E_j | C_i, I)$.


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$$ightarrow P(\boldsymbol{x}_C \,|\, \boldsymbol{x}_E, \Lambda, I)$$

 $\Rightarrow P(\boldsymbol{x}_C | \boldsymbol{x}_E, I) = \int P(\boldsymbol{x}_C | \boldsymbol{x}_E, \Lambda, I) f(\Lambda | I) d\Lambda \quad [by MC!]$



• Bayes theorem:

 $P(oldsymbol{x}_C \,|\, oldsymbol{x}_E, \, \Lambda, \, I) \; \propto \; P(oldsymbol{x}_E \,|\, oldsymbol{x}_C, \, \Lambda, \, I) \cdot P(oldsymbol{x}_C \,|\, I) \,.$



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Indifference w.r.t. all possible spectra

$$P(oldsymbol{x}_C \,|\, oldsymbol{x}_E, \, \Lambda, \, I) ~~ \propto ~~ P(oldsymbol{x}_E \,|\, oldsymbol{x}_C, \, \Lambda, \, I)$$

$$P(\boldsymbol{x}_E \mid x_{C_i}, \Lambda, I)$$



Given a certain number of events in a cause-bin $x(C_i)$, the number of events in the effect-bins, included the 'trash' one, is described by a multinomial distribution:

$$\boldsymbol{x}_E|_{x(C_i)} \sim \operatorname{Mult}[x(C_i), \boldsymbol{\lambda}_i],$$

with

$$\lambda_{i} = \{\lambda_{1,i}, \lambda_{2,i}, \dots, \lambda_{n_{E}+1,i}\} \\ = \{P(E_{1} | C_{i}, I), P(E_{2} | C_{i}, I), \dots, P(E_{n_{E}+1,i} | C_{i}, I)\}$$

 $P(\boldsymbol{x}_E \mid \boldsymbol{x}_C, \Lambda, I)$



 $m{x}_{E}|_{x(C_i)}$ multinomial random vector, $\Rightarrow m{x}_{E}|_{m{x}(C)}$ sum of several multinomials.

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 \Rightarrow Change strategy

Instead of using the original probability inversion (applied directly) to spectra

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we restart from

 $P(C_i | E_j, I) \propto P(E_j | C_i, I) \cdot P(C_i | I).$

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- 1. the sharing of observed events among the cause bins needs to be performed 'by hand';
- 2. a uniform prior $P(C_i | I) = k$ does not mean indifference over all possible spectra.
 - $\Rightarrow P(C_i | I) = k \text{ is a well precise spectrum}$ (in most cases far from the physical one)
 - \Rightarrow VERY STRONG prior that biases the result!

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Oľ

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$$\begin{aligned} x(C_i)|_{x(E_j)} &\approx P(C_i | E_j, I) \cdot x(E_j) \\ x(C_i)|_{\boldsymbol{x}_E} &\approx \sum_{j=1}^{n_E} P(C_i | E_j, I) \cdot x(E_j) \\ x(C_i) &\approx \frac{1}{\epsilon_i} x(C_i)|_{\boldsymbol{x}_E} , \end{aligned}$$

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with $\epsilon_i = \sum_{j=1}^{n_E} P(E_j | C_i, I)$ 4. [*] Uncertainty by 'standard error propagation'

1. λ_i : having each element λ_{ji} the meaning of " p_j " of a Multinomial distribution, their distribution can easily (and conveniently and realistically) modelled by a Dirichlet:

$$\boldsymbol{\lambda}_i ~\sim~ \mathsf{Dir}[oldsymbol{lpha}_{prior} + \left. oldsymbol{x}_E^{MC}
ight|_{x(C_i)^{MC}}] \, ,$$

(The Dirichlet is the prior conjugate of the Multinomial)

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2. uncertainty on λ_i : taken into account by sampling \Rightarrow equivalent to integration

$$\Rightarrow P(\boldsymbol{x}_C | \boldsymbol{x}_E, I) = \int P(\boldsymbol{x}_C | \boldsymbol{x}_E, \Lambda, I) f(\Lambda | I) \, \mathrm{d}\Lambda$$

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4. $x(E_j) \rightarrow \mu_j$: what needs to be shared is not the observed number $x(E_j)$, but rather the estimated true value μ_j : remember $x(E_j) \sim \text{Poisson}[\mu_j]$

$$\mu_j \sim \text{Gamma}[c_j + x(E_j), r_j + 1],$$

Gamma is prior conjugate of Poisson)

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$$x(E_j) \rightarrow \mu_j$$
:

$$\mu_j \sim \operatorname{Gamma}[c_j + x(E_j), r_j + 1],$$

BUT μ_i is real, while the the number of event parameter of a multinomial must be integer \Rightarrow solved with interpolation

5. uncertainty on μ_i : taken into account by sampling

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- $\begin{array}{l} \Rightarrow \text{ problem worked around by ITERATIONS} \\ \Rightarrow \text{ posterior becomes prior of next iteration} \\ \Rightarrow \textbf{Usque tandem?} \end{array}$
 - Empirical approach (with help of simulation):
 - 'True spectrum' recovered in a couple of steps
 - Then the solution starts to diverge towards a wildy oscillating spectrum (any unavoidable fluctuation is believed more and more...)
 - \Rightarrow find empirically an optimum

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 - **regularization** (a subject by itself) my preferred approach
 - regularize the posterior before using as next prior
 - $^{\circ}$ intermediate smoothing \Rightarrow we belief physics is 'smooth'
 - $^\circ\,\ldots$ but 'irregularities' of the data are not washed out (\Rightarrow unfolding Vs parametric inference)

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 - **regularization** (a subject by itself) my preferred approach
 - regularize the posterior before using as next prior
 - \Rightarrow Good compromize and good results
 - \Rightarrow Very 'Bayesian'
 - \Rightarrow No oscillations for $n_{steps} \rightarrow \infty$

Examples

smearing matrix (from 1995 NIM paper)



quite bad! (real cases are usually more gentle)

Examples

smearing matrix (from 1995 NIM paper)



 \Rightarrow watch DEMO

Conclusions

left to users

- 1. "non chiedere all'oste com'è il vino"...
- 2. if I knew how (and was able) to do it better, I had already done it...

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- "iterative" put within parentheses in title (motivated by Zech' classification of methods)
 (a) the spirit of the method is Bayesian
 - (b) the iteration issue is secondary

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- 2. An interesting book: (thanks to Blobel)
 - J. Kaipio and E. Somersalo Statistical and Computational Inverse Problems Springer, 2004

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- 3. Uncertainty due the possible choice among several smearing models, Λ_1 , Λ_2 , etc. (triggered by Marisa Sandhoff's talk)
 - the θ_i sampling can be done at random form either matrix,
 - with weights depending on our beliefs in the different unfolding models
 - (obviously not yet implemented in the R code, and I am not sure I will do it, but it can be implemented in C/C++ versions)

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GdA, NIM A362 (1995) 487

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 \Rightarrow just a honest statement: what is wrong with it?

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Buon divertimento!