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"All models are wrong, but some are useful" (G. Cox)

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'Observed' spectra are often distorted for 'the reasons we know'



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Our aim

from the observed one try to get what we could have got with an ideal detector

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Our aim

from the observed one try to get what we could have got with an ideal detector

But, obviously, not this cyan histogram from this magenta one!

'Observed' spectra are often distorted for 'the reasons we know'



Our aim

from the observed one try to get what we could have got with an ideal detector

But, obviously, not this cyan histogram from this magenta one! \Rightarrow We have to deal with uncertainty and probability

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Although other 'methods' might be more fashionable



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Although other 'methods' might be more fashionable



[Plus other **prescriptions** you might imagine...]

Let's start





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Two-photon invariant mass

ATLAS Experiment at LHC (CERN, Geneva)



ATLAS Experiment at LHC [length: 46 m; Ø 25 m]



pprox 7000 tonnes

pprox 100 millions electronic channels

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Two flashes of 'light' (2 γ 's) in a 'noisy' environment.



Two flashes of 'light' (2 γ 's) in a 'noisy' environment. Higgs $\rightarrow \gamma \gamma$?



Two flashes of 'light' (2 γ 's) in a 'noisy' environment. Higgs $\rightarrow \gamma \gamma$? Probably not...







Quite indirect measurements of something we do not "see"!

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But, can we see our mass?





... or a voltage?





... or our blood pressure?



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Certainly not!



Certainly not!

- ... although for some quantities we can have
- a 'vivid impression' (in the David Hume's sense)



Measuring a mass on a scale



Equilibrium ('physical principle of the measurement'):

 $mg - k\Delta x = 0$ $\Delta x \rightarrow \theta \rightarrow \text{scale reading}$

(with 'g' gravitational acceleration; 'k' spring constant.)

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From the reading to the value of the mass:

scale reading $\xrightarrow{given g, k, "etc."...} m$

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scale reading $\xrightarrow{given g, k, "etc."...} m$



scale reading
$$\xrightarrow{given g, k, "etc."...} m$$

Dependence on 'g': $g \stackrel{?}{=} \frac{GM_{t}}{R_{t}^2}$

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... not even ellipsoidal...



- Earth not spherical...
- ... not even ellipsoidal...
- ...and not even homogeneous.



- Position is usually <u>not</u> at "R_b" from the Earth center;
- Earth not spherical...
- ... not even ellipsoidal...
- ...and not even homogeneous.
- Moreover we have to consider centrifugal effects



Position is usually <u>not</u> at "R₅" from the Earth center;

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- ...and even the effect from the Moon



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▶ ...



left to your imagination...



 $\Delta x \rightarrow \theta \rightarrow$ scale reading:

left to your imagination...

- + randomic effects:
 - stopping position of damped oscillation;
 - variability of all quantities of influence (in the ISO-GUM sense);
 - reading of analog scale.



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 m??
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$\mathsf{Mass} \longrightarrow \mathsf{Reading}$



$\mathsf{Mass} \longrightarrow \mathsf{Reading}$







$\mathsf{Reading} \longrightarrow `\mathsf{true'} \mathsf{ mass}$



$\mathsf{Reading} \longrightarrow \mathsf{`true'} \mathsf{ mass}$



Measurement is nothing but

inferring a model parameter

Physical world \leftrightarrow Science (from Latin '*scio*' – to know)

An Einstein's quote (from his Autobiography) might help:

"Physical concepts are free creations of the human mind, and are not, however it may seem, uniquely determined by the external world" Physical world \leftrightarrow Science (from Latin '*scio*' – to know)

An Einstein's quote (from his Autobiography) might help:

"Physical concepts are free creations of the human mind, and are not, however it may seem, uniquely determined by the external world"

And, again there, referring to his revolutianary ideas:

"The type of critical reasoning which was required for the discovery of this central point was decisively furthered, in my case, especially by the reading of David Hume's and Ernst Mach's philosophical writings"



The observed 'data' is certain:











The <u>observed</u> 'data' is certain: \rightarrow 'true value' uncertain "Data uncertainty"? **???** Data corrupted? Even if the data were corrupted, the <u>data</u> were the corrupted data**!!**...



"Data uncertainty"? **???** Data corrupted? Even if the data were corrupted, the <u>data</u> were the corrupted data!!...

[Unless we are talking of 'future data' or of 'somebody else data' ...]

Observation \rightarrow value of a quantity



scale reading given g, k, "etc."... m



Observation \rightarrow value of a quantity



scale reading $\xrightarrow{}$ given g, k, "etc."... m

 \Rightarrow Measurement is not simply 'reading' a value on an instrument



Observation \rightarrow value of a quantity



scale reading
$$\xrightarrow{given g, k, "etc."...} m$$

⇒ Measurement is not simply 'reading' a value on an instrument (A reading, without proper contextualization, is ... just a number)

$Observation \rightarrow value \ of \ a \ quantity$



scale reading
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⇒ Measurement is not simply 'reading' a value on an instrument (A reading, without proper contextualization, is ... just a number)

Mistrust the

"dogma of the Immaculate Observation"!

$Observations \rightarrow hypotheses$

This problem occurs not only "determining" *the* value of a physical quantity.

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• Experimental observation ('data') \rightarrow responsible cause.



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• Experimental observation ('data') \rightarrow responsible cause.

(But logically no substantial difference.)

$\mathsf{Causes} \to \mathsf{effects}$

The same apparent cause might produce several, different effects



Given an observed effect, we are not sure about the exact cause that has produced it.

$\mathsf{Causes} \to \mathsf{effects}$

The same *apparent* cause might produce several, different effects



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$\mathsf{Causes} \to \mathsf{effects}$

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Given an observed effect, we are not sure about the exact cause that has produced it.

 $\textbf{E_2} \Rightarrow \{\textit{C}_1, \textit{C}_2, \textit{C}_3\}?$



The "essential problem" of the Sciences

"Now, these problems are classified as probability of causes, and are most interesting of all for their scientific applications.


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I play with a gentleman whom I do not know. He has dealt ten times, and he has turned the king up six times. What is the chance that he is a sharper?



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(H. Poincaré – Science and Hypothesis)

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Why we (or most of us) have not been taught how to tackle this kind of problems?

Basic rules of probability $\sqrt{}$

1.
$$0 \leq P(A \mid I) \leq 1$$

$$2. \qquad P(\Omega \mid I) = 1$$

3.
$$P(A \cup B \mid I) = P(A \mid I) + P(B \mid I)$$
 [if $P(A \cap B \mid I) = \emptyset$]

4.
$$P(A \cap B | I) = P(A | B, I) \cdot P(B | I) = P(B | A, I) \cdot P(A | I)$$

Remember that probability is always conditional probability!

I is the background condition (related to information ${}^{\prime}I'_{s}$) \rightarrow usually implicit (we only care about 're-conditioning')

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Remember that probability is always conditional probability!

- I is the background condition (related to information I_s')
- $\rightarrow\,$ usually implicit (we only care about 're-conditioning')
- Note: 4. <u>does not</u> define conditional probability. (Probability is <u>always</u> conditional probability!)
 - ⇒ easily extended to uncertain numbers ('random variables')

"Since the knowledge may be different with different persons

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Probability depends on the status of information of the *subject* who evaluates it.



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"Thus whenever we speak loosely of 'the probability of an event', it is always to be understood: probability with regard to a certain given state of knowledge"

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$P(E) \longrightarrow P(E \mid I_s(t))$

where $l_s(t)$ is the information available to subject s at time t.

Mathematics of beliefs

An even better news:

The fourth basic rule can be fully exploited!



Mathematics of beliefs

An even better news:

The fourth basic rule can be fully exploited!

(Liberated by a curious ideology that forbids its use)



A simple, powerful formula

P(A|B|I)P(B|I) = P(B|A,I)P(A|I) $P(A|B) = \frac{P(B|A)P(A)}{P(B)}$

A simple, powerful formula

Take the courage to use it!

 $P(A|B) = \frac{P(B|A) P(A)}{P(B)}$

A simple, powerful formula

$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \mathbb{P}(A)}{\mathbb{P}(B)}$ It's easy if you try...! (c) GdA, Roma 26/02/24, 24/69

"The greater the probability of an observed event given any one of a number of causes to which that event may be attributed, the greater the likelihood of that cause {given that event}.

$P(C_i \mid E) \propto P(E \mid C_i)$

"The greater the probability of an observed event given any one of a number of causes to which that event may be attributed, the greater the likelihood of that cause {given that event}. The probability of the existence of any one of these causes {given the event} is thus a fraction whose numerator is the probability of the event given the cause, and whose denominator is the sum of similar probabilities, summed over all causes.

$$P(C_i \mid E) = \frac{P(E \mid C_i)}{\sum_j P(E \mid C_j)}$$

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$$P(C_i | E) = \frac{P(E | C_i) P(C_i)}{P(E)}$$

(Philosophical Essai on Probabilities)

[In general $P(E) = \sum_{j} P(E | C_j) P(C_j)$ (weighted average, with weigths being the probabilities of the conditions) if C_j form a complete class of hypotheses]

$$P(C_i | E) = \frac{P(E | C_i) P(C_i)}{P(E)} = \frac{P(E | C_i) P(C_i)}{\sum_j P(E | C_j) P(C_j)}$$

"This is the fundamental principle ^(*) of that branch of the analysis of chance that consists of reasoning a posteriori from events to causes"

(*) In his "Philosophical essay" Laplace calls 'principles' the 'fundamental rules'.

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Note: denominator is just a normalization factor.

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Most convenient way to remember Bayes theorem

Telling it with Gauss' words

A quote from the Princeps Mathematicorum (Prince of Mathematicians) is <u>a must</u>.



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$$P(C_i | data) = \frac{P(data | C_i)}{P(data)} P_0(C_i)$$



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"post illa observationes" "ante illa observationes" (Gauss)



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Arguments used to derive 'his' error function

- 1. $f(\mu | \{x\}) \propto f(\{x\} | \mu) \cdot f_0(\mu)$
- f₀(μ) 'flat': all values 'a priori' equally possible ("... aeque probabilia fuisse")
- 3. posterior maximized at $\mu = \overline{x}$

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Indeed Gauss had clear ideas about the role of the priors and also of the fact the, strictly speaking, the 'Gaussian' "cannot express the probability of the errors" (!!) [cfr e.g. arXiv:2003.10878] ("All models are wrong...") © GdA, Roma 26/02/24, 27/69

Personal recollections



(Thanks for patience and compassion...)

© GdA, Roma 26/02/24, 28/69
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- Astonished by how sensible the results were!
- Later I discovered the Bayesian world.

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\Rightarrow Short reminder

HERA Physics



Main physics goal: proton structure (+...., including searches)

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Deep inelastic scattering at HERA



[OVERSIMPLIFIED^(*) diagram, from CERN Courier, 2015/08] (*)A shame: it confirms my adversion to popularization of Science

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A proton is a complicated 'structured' dynamical object...



A proton is a complicated 'structured' dynamical object...







electron in the final state (+ hadronic jet from scattered q);





 electron in the final state (+ hadronic jet from scattered q);
 kinematical quantities x and Q² 'easily' measured (in principle ...).









final state neutrino not directly observable;
 x and Q² have to be measured only from ('current') jet hadrons...



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many of which are lost in the beam pipe, depending on x and Q².



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many of which are lost in the beam pipe, depending on x and Q².

• Measurement of x and Q^2 becomes challenging!!



VERY HARD scattering (rare event)

Esempio de migrezioni de 4 dx, dQ2 } a seconda dei mebodi de nicostrazione



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How to go back from reconstructed quantities to physical quantities?

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Note: 2D smearing: \rightarrow 2D unfolding!

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How to go back from reconstructed quantities to physical quantities?

Note: 2D smearing: \rightarrow 2D unfolding! \rightarrow **Multidimensional unfolding**

The basic idea — 1D for clarity


The basic idea — 1D for clarity



How to use if for unfolding? ells in => Pullidiments

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• 'Cause bin' $C_i \rightarrow$ 'Effect bin' E_j



Cause bin' C_i → 'Effect bin' E_j
 The C_i can be defined in any space:
 ⇒ multidimensional



- 'Cause bin' $C_i \rightarrow$ 'Effect bin' E_j
- ► The C_i can be defined in any space: ⇒ multidimensional
- Inefficiencies are described by adding to the effect cells a Trash cell (it will be clearer in a while)



Background can also be naturally included, by just adding an extra cause bin:

- obviously we have to 'know' (by MC) how it will contribute (see figure in the previous slide)
- also several sources of background can be included.

Sharing the observed events: 1. evaluate $P(C_i|E_i)$ => Tullidiment "smeaning maker" $P(c:|E_j) = \frac{P(E_j|c_i) \cdot P_o(c_i)}{\sum_{e=1}^{n_c} P(E_j|c_e) \cdot P_o(c_e)}$ · Po((i) - initial probabilities (Z. Po((:) = 1) - wich? best guess - if P(Ck)=0 it will never be updated: [inservicion of delador.] · ¿P(E; 1Ci)=1 each effect must come from one of the causes (eventually also background © GdA, Roma 26/02/24, 42/69

Sharing the observed events: 2. evaluate $n(E_i) \rightarrow n(C_i)$

$$\varepsilon_{i} = \sum_{j=1}^{n_{E}} P(E_{j}|C_{i}) \qquad o \leq \varepsilon_{i} \leq 1 \qquad efficiency for C_{i}$$

$$\varepsilon_{i} \sim o, \text{ hic } \text{ mark leaves} \text{ to be observed be one} \qquad of the worde E_{j}$$

$$N_{obs} \qquad : \qquad \underline{n}(E) = \{n(E_{1}), \dots, n(E_{n_{E}})\}$$

$$n(c_{i}) \qquad n(C_{2}) \qquad n(c_{n})$$

$$\hat{n}(C_{i})|_{observed} \approx \sum_{j=1}^{n_{E}} n(E_{j}) \cdot P(C_{i}|E_{j})$$

$$\frac{n(C_{i})}{2k}|_{observed} \qquad (619)$$

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$$n(c_{i}) \qquad n(\varepsilon_{2}) \qquad n(c_{n})$$

$$\widehat{n}(c_{i})|_{observed} = \sum_{j=1}^{n_{E}} n(\varepsilon_{j}) \cdot P(C_{i} | \varepsilon_{j})$$

$$-225 - \qquad (1)$$$$

... also taking into account the inefficiencies (events which went to 'Trash')

$$\hat{n}(c;) = \frac{1}{\varepsilon_i} \sum_{j=i}^{\varepsilon_i} n(\varepsilon_j) \cdot P(c; (\varepsilon_j))$$

At this point you might have several doubts and questions (e.g. about priors) \rightarrow Please wait a while.

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 - \rightarrow 'NIM' A362 (1995) 487

Example of 2D unfolding (from NIM paper)



A multidimensional unfolding method based on Bayes' theorem

G. D'Agostini *

Università "La Sapienza" and INFN, Roma, Italy

Received 5 August 1994; revised form received 2 March 1995

Abstract

Bayes' theorem offers a natural way to unfold experimental distributions in order to get the best estimates of the true ones. The weak point of the Bayes approach, namely the need of the knowledge of the initial distribution, can be overcome by an iterative procedure. Since the method proposed here does not make use of continuous variables, but simply of cells in the spaces of the true and of the measured quantities, it can be applied in multidimensional problems.

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 - STRONG POINT, because it allows to combine new pieces of information with prior knowledge.
 - It is an ESSENTIAL INGREDIENT of probability theory if we want to reason from effects to causes (see e.g. Laplace and Gauss).

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A kind of technical recap

now that basic ideas shoud be received

(focusing on 1D unfolding, with life examples)



Discretized unfolding





Discretized unfolding



Background ('known') is just a an extra cell

 $n_C \rightarrow n_C + 1$

(Hereafter just included among the causes - several BG's can be included)

Vectors of interest:

Discretized unfolding



Background ('known') is just a an extra cell

$$n_C \rightarrow n_C + 1$$

(Hereafter just included among the causes - several BG's can be included)

Vectors of interest:

x_C: true spectrum (nr of events in cause bins)

x_E: observed spectrum (nr of events in effect bins)

Remark: parametric inference Vs unfolding

 $f(\theta | \{x, y\})$:

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Remark: parametric inference Vs unfolding

$f(\boldsymbol{\theta} | \{\boldsymbol{x}, \boldsymbol{y}\})$:

- $\rightarrow\,$ probabilistic parametric inference
 - \Rightarrow it relies on the kind of functions parametrized by ${\boldsymbol \theta}$

$$oldsymbol{\mu}_{_{\mathcal{Y}}} = oldsymbol{\mu}_{_{\mathcal{Y}}}(oldsymbol{\mu}_{_{\mathcal{X}}};oldsymbol{ heta})$$
 with (because of errors)

$$\begin{array}{cccc} \mu_{\mathsf{x}_i} & \longrightarrow & \mathsf{x}_i \\ \mu_{\mathsf{y}_i} & \longrightarrow & \mathsf{y}_i \end{array}$$

Remark: parametric inference Vs unfolding

$f(\boldsymbol{\theta} | \{\boldsymbol{x}, \boldsymbol{y}\})$:

→ probabilistic parametric inference ⇒ it relies on the kind of functions parametrized by θ $\mu_y = \mu_y(\mu_x; \theta)$ with (because of errors)

$$\begin{array}{cccc} \mu_{x_i} & \longrightarrow & x_i \\ \mu_{y_i} & \longrightarrow & y_i \end{array}$$

 \Rightarrow data distilled into θ : $\rightarrow f(\theta | data)$

Remark: parametric inference Vs unfolding *f*(θ | {x, y}):

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BUT sometimes we wish to interpret the data as little as possible

⇒ just public 'something equivalent' to an experimental distribution, with the bin contents fluctuating according to an underlying multinomial distribution, but having possibly got rid of physical and instrumental distortions, as well as of background.
Remark: parametric inference Vs unfolding *f*(θ | {x, y}):

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$$\boldsymbol{\mu}_{y} = \boldsymbol{\mu}_{y}(\boldsymbol{\mu}_{x};\boldsymbol{\theta})$$

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BUT sometimes we wish to interpret the data as little as possible

- ⇒ just public 'something equivalent' to an experimental distribution, with the bin contents fluctuating according to an underlying multinomial distribution, but having possibly got rid of physical and instrumental distortions, as well as of background.
- ⇒ Unfolding (deconvolution)

Invert smearing matrix?



Invert smearing matrix? In general is a bad idea: not a rotational problem <u>but</u> an inferential problem!



$$\begin{array}{l} \text{Imagine } S = \begin{pmatrix} 0.8 & 0.2 \\ 0.2 & 0.8 \end{pmatrix}; \rightarrow U = S^{-1} = \begin{pmatrix} 1.33 & -0.33 \\ -0.33 & 1.33 \end{pmatrix} \\ \text{Let the true be } s_t = \begin{pmatrix} 10 \\ 0 \end{pmatrix}; \rightarrow s_m = S \cdot s_t = \begin{pmatrix} 8 \\ 2 \end{pmatrix}; \\ \text{If we measure } s_m = \begin{pmatrix} 8 \\ 2 \end{pmatrix} \rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 10 \\ 0 \end{pmatrix} \sqrt{2} \end{array}$$

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Imagine $S = \begin{pmatrix} 0.8 & 0.2 \\ 0.2 & 0.8 \end{pmatrix}$: $\rightarrow U = S^{-1} = \begin{pmatrix} 1.33 & -0.33 \\ -0.33 & 1.33 \end{pmatrix}$ Let the true be $s_t = \begin{pmatrix} 10 \\ 0 \end{pmatrix}$: $\rightarrow s_m = S \cdot s_t = \begin{pmatrix} 8 \\ 2 \end{pmatrix}$; If we measure $s_m = \begin{pmatrix} 8 \\ 2 \end{pmatrix} \rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 10 \\ 0 \end{pmatrix} \sqrt{2}$ BUT if we had measured $\begin{pmatrix} 9\\1 \end{pmatrix} \rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 11.7\\-1.7 \end{pmatrix}$ if we had measured $\begin{pmatrix} 10 \\ 0 \end{pmatrix} \rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 13.3 \\ -3.3 \end{pmatrix}$

$$\begin{array}{l} \text{Imagine } S = \begin{pmatrix} 0.8 & 0.2 \\ 0.2 & 0.8 \end{pmatrix} : \rightarrow U = S^{-1} = \begin{pmatrix} 1.33 & -0.33 \\ -0.33 & 1.33 \end{pmatrix} \\ \text{Let the true be } s_t = \begin{pmatrix} 10 \\ 0 \end{pmatrix} : \rightarrow s_m = S \cdot s_t = \begin{pmatrix} 8 \\ 2 \end{pmatrix} ; \\ \text{If we measure } s_m = \begin{pmatrix} 8 \\ 2 \end{pmatrix} \Rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 10 \\ 0 \end{pmatrix} \checkmark \\ \begin{array}{l} \text{BUT} \\ \text{if we had measured } \begin{pmatrix} 9 \\ 1 \end{pmatrix} \Rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 11.7 \\ -1.7 \end{pmatrix} \\ \text{if we had measured } \begin{pmatrix} 10 \\ 0 \end{pmatrix} \Rightarrow S^{-1} \cdot s_m = \begin{pmatrix} 13.3 \\ -3.3 \end{pmatrix} \end{array}$$

Indeed, matrix inversion is recognized to producing 'crazy spectra' and even negative values (unless there are so large numbers of events in bins such that fluctuations around expectations are negligeable)

Bin to bin analysis?

En passant:

► OK if the are no migrations: → each bin is an 'independent issue', treated with a binomial process, given some efficiencies.



Bin to bin analysis?

En passant:

- OK if the are no migrations:
 - \rightarrow each bin is an 'independent issue',
 - treated with a binomial process, given some efficiencies.

Otherwise

- 'error analysis' troublesome
 (just imagine e.g. that a bin has an 'efficiency' > 1,
 because of migrations from other bins);
- iteration is important

(efficiencies depend on 'true distribution')



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 x_C : true spectrum (nr of events in cause bins) x_E : observed spectrum (nr of events in effect bins)



x_C: true spectrum (nr of events in cause bins)

 x_E : observed spectrum (nr of events in effect bins) Our aim:

- not to find <u>the</u> true spectrum
- but, more modestly, <u>rank in beliefs</u> all possible spectra that might have caused the observed one:
 ⇒ P(x_C | x_E, l)



► $P(\mathbf{x}_C | \mathbf{x}_E, I)$ depends on the knowledge of *smearing matrix* Λ , with $\lambda_{ji} \equiv P(E_j | C_i, I)$.



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▶ but A is itself uncertain, because inferred from MC simulation: ⇒f(A | I)



- ► $P(\mathbf{x}_C | \mathbf{x}_E, I)$ depends on the knowledge of *smearing matrix* Λ , with $\lambda_{ji} \equiv P(E_j | C_i, I)$.
- but ∧ is itself uncertain, because inferred from MC simulation: ⇒f(∧ | I)
- for each possible Λ we have a pdf of spectra: $\rightarrow P(\mathbf{x}_C \mid \mathbf{x}_E, \Lambda, I)$



- ► $P(\mathbf{x}_C | \mathbf{x}_E, I)$ depends on the knowledge of *smearing matrix* Λ , with $\lambda_{ji} \equiv P(E_j | C_i, I)$.
- ▶ but Λ is itself uncertain, because inferred from MC simulation: $\Rightarrow f(\Lambda | I)$

► for each possible Λ we have a pdf of spectra: $\rightarrow P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I)$ $\Rightarrow P(\mathbf{x}_C | \mathbf{x}_E, I) = \int P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) f(\Lambda | I) d\Lambda$ [by MC!]





 $P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) \propto P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I) \cdot P(\mathbf{x}_C | I).$





Bayes theorem:

 $P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) \propto P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I) \cdot P(\mathbf{x}_C | I).$

Indifference w.r.t. all possible spectra

$$P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) \propto P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I)$$



Given a certain number of events in a cause-bin $x(C_i)$, the number of events in the effect-bins, included the 'trash' one, is described by a multinomial distribution:

$$\mathbf{x}_{E}|_{\mathbf{x}(C_{i})} \sim \operatorname{Mult}[\mathbf{x}(C_{i}), \boldsymbol{\lambda}_{i}],$$

with

$$\lambda_{i} = \{\lambda_{1,i}, \lambda_{2,i}, \dots, \lambda_{n_{E}+1,i}\} \\ = \{P(E_{1} | C_{i}, I), P(E_{2} | C_{i}, I), \dots, P(E_{n_{E}+1,i} | C_{i}, I)\}$$

$P(\boldsymbol{x}_E \mid \boldsymbol{x}_C, \Lambda, I)$



 $\begin{aligned} \mathbf{x}_{E}|_{\mathbf{x}(C_{i})} & \text{multinomial random vector,} \\ \Rightarrow \mathbf{x}_{E}|_{\mathbf{x}(C)} & \text{sum of several multinomials.} \end{aligned}$



 $P(\mathbf{x}_F | \mathbf{x}_C, \Lambda, I)$



 $x_E|_{x(C_i)}$ multinomial random vector, $\Rightarrow x_E|_{x(C)}$ sum of several multinomials. BUT

no 'easy' expression for $P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I)$

 $P(\mathbf{x}_F | \mathbf{x}_C, \Lambda, I)$



 $x_E|_{x(C_i)}$ multinomial random vector, $\Rightarrow x_E|_{x(C)}$ sum of several multinomials. BUT

no 'easy' expression for $P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I)$ \Rightarrow STUCK!

$P(\mathbf{x}_E \,|\, \mathbf{x}_C, \, \Lambda, \, I)$



 $x_E|_{x(C_i)}$ multinomial random vector, $\Rightarrow x_E|_{x(C)}$ sum of several multinomials. BUT

no 'easy' expression for $P(\mathbf{x}_E \mid \mathbf{x}_C, \Lambda, I)$

- \Rightarrow STUCK!
- \Rightarrow Change strategy

Instead of using the original probability inversion (applied directly) to spectra

$$P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) \propto P(\mathbf{x}_E | \mathbf{x}_C, \Lambda, I) \cdot P(\mathbf{x}_C | I),$$

we restart from

 $P(C_i | E_j, I) \propto P(E_j | C_i, I) \cdot P(C_i | I).$



The rescue trick

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$$P(C_i | E_j, I) = \frac{P(E_j | C_i, I) \cdot P(C_i | I)}{\sum_i P(E_j | C_i, I) \cdot P(C_i | I)},$$

or

$$\theta_{ij} = \frac{\lambda_{ji} \cdot P(C_i \mid I)}{\sum_i \lambda_{ji} \cdot P(C_i \mid I)},$$

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$$\begin{aligned} x(C_i)|_{x(E_j)} &\approx P(C_i | E_j, I) \cdot x(E_j) \\ x(C_i)|_{\boldsymbol{X}_E} &\approx \sum_{j=1}^{n_E} P(C_i | E_j, I) \cdot x(E_j) \\ x(C_i) &\approx \frac{1}{\epsilon_i} | x(C_i)|_{\boldsymbol{X}_E} , \end{aligned}$$

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4. [*] Uncertainty by 'standard error propagation'

1. λ_i : having each element λ_{ji} the meaning of " p_j " of a Multinomial distribution, their distribution can easily (and conveniently and realistically) modelled by a Dirichlet:

$$\lambda_i \sim \text{Dir}[lpha_{prior} + \mathbf{x}_E^{MC}\Big|_{x(C_i)^{MC}}],$$

(The Dirichlet is the prior conjugate of the Multinomial)

1.
$$\lambda_i$$
:
 $\lambda_i \sim \operatorname{Dir}[\alpha_{prior} + \mathbf{x}_E^{MC}\Big|_{x(C_i)^{MC}}],$

2. uncertainty on λ_i :

taken into account by sampling \Rightarrow equivalent to integration

$$\Rightarrow P(\mathbf{x}_C | \mathbf{x}_E, I) = \int P(\mathbf{x}_C | \mathbf{x}_E, \Lambda, I) f(\Lambda | I) d\Lambda$$



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$$\mathbf{x}_{C}|_{x(E_{j})} \sim \operatorname{Mult}[x(E_{j}), \theta_{j}],$$

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$$\mathbf{x}_C|_{\mathbf{x}(E_i)} \sim \operatorname{Mult}[\mathbf{x}(E_j), \boldsymbol{\theta}_j],$$

 x(E_j) → μ_j: what needs to be shared is not the observed number x(E_j), but rather the estimated true value μ_j: remember x(E_j) ~ Poisson[μ_j]

$$\mu_j \sim \operatorname{Gamma}[c_j + x(E_j), r_j + 1],$$

(Gamma is prior conjugate of Poisson)

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4. $x(E_j) \rightarrow \mu_j$:

$$\mu_j ~\sim~ \operatorname{Gamma}[c_j + x(E_j), r_j + 1],$$

BUT μ_i is real, while the the number of event parameter of a multinomial must be integer \Rightarrow solved with interpolation

5. uncertainty on μ_i : taken into account by sampling

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Empirical approach (with help of simulation):

- 'True spectrum' recovered in a couple of steps
- Then the solution starts to diverge towards a wildy oscillating spectrum (any unavoidable fluctuation is believed more and more...)

 \Rightarrow find empirically an optimum

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 - ► intermediate smoothing ⇒ we belief physics is 'smooth'
 - ... but 'irregularities' of the data are not washed out

 $(\Rightarrow$ unfolding Vs parametric inference)

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 regularization (a subject by itself) my preferred approach

- regularize the posterior before using as next prior
- \Rightarrow Good compromize and good results
- ⇒ Very 'Bayesian'
- \Rightarrow No oscillations for $\mathit{n_{steps}} \rightarrow \infty$

Examples

Smearing matrices (from 1995 NIM paper)



quite bad! (real cases are usually more gentle)

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\Rightarrow watch DEMO

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- Bayesian networks;
- MCMC for integration;
- JAGS code to perform the practical work

It is possible to do better? Draft 16 October 2012

It is possible to do better? Draft 16 October 2012

Straight Bayesian unfolding performed by MCMC techniques using JAGS/rjags

G. D'Agostini¹ and E. Franco²

¹ Università "La Sapienza" and INFN, Rome, Italy (giulio.dagostini@roma1.infn.it, http://www.roma1.infn.it/~dagos) ² INFN, Sezione di Roma 1, Rome, Italy (enrico.franco@roma1.infn.it)

Abstract

The observation that the probability 'vector' of the observed bins can be expressed as product of the smearing matrix and the probability 'vector' of the cause bins avoids to formulate the discretized unfolding problem in terms of sum of multinomial distributions. As a result, the joint distribution of all uncertain quantities entering the problem can be factorized by as a product of conditional probabilities, that can be viewed as a Probabilistic ('Bayesian') Graphical Model and solved using modern Markov Chain Monte Carlo methods, finally avoiding the 'dirty trick' of the iterations. In particular, we have implemented these ideas in the JAGS program, run under R using rjags

Results on the same toy models (Oct 2012) - No iterations!



Results on the same toy models (Oct 2012) - No iterations!



GdA, Roma 26/02/24, 63/69

Stay tuned!



The End

For references, also related to the discussion:

Lectures to Phd students (lickable link)



Linked notes



$p = \frac{\# \text{ favorable cases}}{\# \text{ possible equiprobable cases}}$

$$p = \frac{\# \text{ favorable cases}}{\# \text{ possible equiprobable cases}} \Rightarrow \text{Loop!}$$



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Cheating students...

 $p = \frac{\# \text{ favorable cases}}{\# \text{ possible equally possible cases}}$

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But for me it was a serious shock that induced me to rithink the probabilistic issues







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- ▶ Having in hand a faulty item, evaluate the probability that it was produced by each of the three machines: $\rightarrow P(M_i | \text{faulty item})$

Go back

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Some contributions on the subject:

- Jeffreys Priors versus Experienced Physicist Priors Arguments against Objective Bayesian Theory, arXiv:physics/9811045
- Overcoming prior anxiety, arXiv:physics/9906048
- Constraints on the Higgs Boson Mass from Direct Searches and Precision Measurements [with G. Degrassi], arXiv:hep-ph/9902226
- Inferring the intensity of Poisson processes at limit of detector sensitivity (with a case study on gravitational wave burst search) [with P. Astone], arXiv:hep-ex/9909047
- ▶ Bayesian reasoning in data analysis A critical introduction, 2003